

On Student's- t ARMA Modelling of Missing Values

ABSTRACT

In this paper we intend to mitigate the missing data problem in the context of univariate ARMA time series model. Our main objective was to derive imputation estimators for ARMA models under the student's- t distributional assumptions and evaluate their imputation performance. We utilized the method of optimal interpolation criterion of missing values. We carried out a simulation of 1000 samples using statistical R software. We carried out imputation of missing values using the derived estimators for ARMA (1,1) and ARMA (1,2). We observed that ARMA (1,2) did predicted the missing values better than ARMA (1,1). We again carried out a comparison of the derived imputation estimators of time series with the convectional imputation techniques of missing values. They included K^{th} - Nearest Neighborhood (KNN), Artificial Neural Networks (ANN) and Kalman filters. The imputation metrics were calculated to compare the above techniques. The results obtained from the comparison of the simulation study were that the ANN, KNN and the Kalman filters were the best in predicting the missing values. The proposed models of ARMA also did competed well but not as the first three techniques.

Keywords: Autoregressive Moving Average model, Imputation, Interpolation, Missing Values, Student - t distribution, Optimal Linear interpolation, Kalman filters, Artificial Neuron Networks

1. INTRODUCTION

An autoregressive moving average or (ARMA) model, is a process of a linear combination of an autoregressive or AR, a moving average or MA and an error or an innovation noise. The (ARMA) process is an important stationary time series model that plays a crucial role in the modeling of time series data. In most cases, the errors' noises for the general (ARMA) models are always assumed to take normal distribution its modelling [1].

Generally, most time series models assume normality in their estimation and predictions. Normality is always taken by these models because of its simplicity and availability of modeling techniques that can handle normality more comfortably [2, 3]. However, there are cases where some practical scenarios may not capture normality. Utilizing normality assumptions to model such cases where normality is not obeyed might not work [4].

Asymmetry and non-normality is an alternative way of handling data that do not obey normality. Asymmetric distribution offers more reliable and robust results when applied to real phenomena with asymmetric, non-normality, skewness and flat tails characteristics [5]. The most utilized distributions to capture asymmetry, flat and long tails, are; Student- t distribution, Normal inverse Gaussian and generalized error distribution just to mention a few. The characteristics of these distributions can be adjusted to adapt data that displays, heavy tails, skewness, non-normality and asymmetric features.

During modelling of time series data, researchers normally encounter with missing data.

Missing values occurs when a value or data values in a given variable are not given for many reasons. Missing values masks and veils trends and patterns in time series analysis. Missing values has far reaching effects on analysis and forecasting associated to trends and patterns for time series data [6, 7]. Time series analysis does experiences this challenge quite often. Therefore, many researchers, have given rise to numerous ways of recouping missing data in various disciplines, including time series.

One of the ways to deals with missing values is through the deletion of missing parts. The deletion is a way of eliminating missing sets or the variables containing missing values by simply deleting the missing data. However, deletion of missing values has a major setback of eliminating vital information, which adversely impact the prediction, forecasts and eventually interfere with the inference on vital information for decision making [8, 9].

Imputation is the only way out to mitigate missing value problems. Imputation is an important step that should not be skipped in any data analysis process. Imputation enhances cleanliness in data for further handling and processing. A sufficient data imputation algorithm must improve on the data productivity, data analysis, data visualization and data investigations [2]. During imputation of missing data, there are various pertinent issues that need to be taken into consideration ensure adequate restoration of missing data. One of the important characteristics that any data should portray during analysis is normality. In practice, normality might not be realized for various reason. To cope up with asymmetry and non-normality, data can be transformed to make it capture normality. However, this step has been dismissed owing to the fact that vital statistical properties of the transformed data may be destroyed. Due to this reason, asymmetric distributions have been proven to capture and work well under non-normality, asymmetry and skewness cases [10]

Time series models like autoregressive moving average (ARMA) models were design in such a way that they only work with complete data. If missing data occurs during modeling using the time series models, then modeling might be hampered. This way, researches has been carried out to develop statistical imputation algorithms and estimators through robust statistical estimation techniques like Likelihood Estimation [11], Least Squares Method [12], Expectation Maximization [13], and Interpolation.

1.1 AUTOGRESSIVE MOVING AVARAGE MODEL

The Autoregressive moving average (ARMA) model is a time series model that is widely used for analyzing time series data. The general form of ARMA (p, q) is given by;

$$k_t = \sum_{i=1}^p \theta_i k_{t-i} + \sum_{j=1}^q \varphi_j e_{t-j} + e_t \quad (1)$$

Where e_t in the literature is assumed to be Gaussian white noise with zero mean and a constant variance δ_{α}^2 . In this context, we will have e_t assuming the Student's t distribution. It is taken that this process should be causal-stationary and invertible so that the roots of the process lie outside the unit circle.

Time series models that has considered the imputation of missing values under Gaussian or normal assumptions are numerous in the literature. For instance, author [14] did proposed an algorithm for filling missing values for ARMA time series model when their innovations consider a stable Gaussian assumption. Authors [15, 16] considered the missing value problem under the AR processes of time series models. Their studies focused on letting their AR error terms assume Gaussian approach. Author [17] considered imputation of missing values with pure bilinear time series models. The innovations for the series in their study,

assumed normal distribution.

Time series models that have considered imputation of missing values under asymmetric assumptions include [18, 19], their study considered a vector autoregressive (VAR) time series, to model data with missing data under Multivariate heavy tailed, Student t distribution approach.

We made an observation, that there is a gap in imputing missing values under time series analysis, especially when their innovations follow asymmetric distributions. This is the gap we wish to pay our attention to. The objective for this paper was to develop some imputation estimators for ARMA time series modeling under Student t -distribution assumptions via optimal linear interpolation.

1.2 THE STUDENT T DISTRIBUTION

The Student t -distribution is a continuous distribution that belongs to the family of continuous probability distributions. It has wide range of applications ranging from statistics and other genres of sciences. This distribution has been used in capturing heavy tails, asymmetry and non-normality. Such attributes of heavy tails and asymmetry are mostly displayed by economic data, business data and finance datasets [20, 21].

The standard student t given distribution is given by;

$$f(x) = \frac{1}{\sqrt{v}\beta \left(\frac{v}{2}, \frac{1}{2}\right)} \left(1 + \frac{x^2}{v}\right)^{-\left(\frac{1+v}{2}\right)}$$

$$-\infty < x < \infty, \text{ where } v > 0 \quad (2)$$

Some known characteristics of the Student t distribution on mean, variance, skewness and kurtosis are given by,

- i. Mean (α_1) = $E(x) = 0$
- ii. Variance
Var (x) = $\beta_2 = \frac{v}{v-2}, v > 2$
- iii. Skewness

$$\gamma_1(x) = \frac{\beta_3}{\beta_2^{\frac{3}{2}}} = 0$$

- iv. Kurtosis

$$\gamma_2(x) = \frac{\beta_4}{\beta_2^2} = \frac{3(v-2)}{v-4}, \quad v > 4$$

A challenge arises when utilizing this distribution to capture asymmetry and heavy tail-ness in presence of missing data. This is because the student t -distribution is crafted to be modeled and utilized under a case when the all data is present.

The student t -distribution is one of the most commonly used heavy tailed distribution in modellings. The authors [22, 21, 22] have considered their time series errors following student's t assumptions for missing value estimations.

1.2 OPTIMAL LINEAR INTERPOLATION OF MISSING VALUES

The proposed method was suggested by author [12]. Also, the above mentioned method has been utilized by author [22] for the imputation of missing values under bilinear time series models. For the purposes of computation of missing values in autoregressive moving average (ARMA) time series models. They explained how the imputation computation is arrived at using the following statements, that;

Suppose an observation k_m is a missing value out of a set of n -possible observations generated by an ARMA(p, q) process. Let the subspace Q_m^* be the allowable space of a linear estimator of k_m based on observed values $k_t, k_{t-1}, \dots, k_{m-1}$ that are given by $Q_m^* = Q_p\{k_t: t \leq n; t \neq m\}$. The projection of k_t on to Q_m^* denoted as $P_{S_m}^{k_m}$ such that the disp $\{K_m - P_{S_m}^{k_m}\}$ is minimized, that is basically the minimum dispersion of the linear estimator. Direct computation of the projection of the K_m on to Q_m^* would be complicated since the subspace $Q_1 = Q_p\{k_{m-1}, k_{m-2}, \dots\}$ and Q_m^* are not independent of each other and thus we consider the evaluation of the projection on to two disjoint subspace of Q_m^* . To achieve this, we express Q_m^* as a direct sum of subspaces Q_1 and another subspace, say Q_* such that $Q_m^* = Q_1 \oplus Q_*$. A possible sub-space is $Q_* = Q_p\{k_i - k_i^t; i > m + 1\}$. Where k_i^t is based on the values $\{k_{m-1}, k_{m-2}, \dots\}$. The existence of subspaces Q_1 and Q_* are shown in the following lemma;

Lemma

Suppose k_t is non-determined stationary process defined on the probability space (Ω, β, ρ) . Then the subspace Q_m^* is the direct sum of subspaces Q_1 and Q_* as defined in the above norm.

Proof

Suppose $K_* \in S_m^*$ then K_* can be represented as;

$$K_* = Z^n + \sum a_i K_i^t = (K + \sum a_i K_i^t) + \sum a_i K_i^t \text{ where } K \in S_1 \quad (3)$$

So clearly, the two components in the above equation (2) are independent. The best linear estimators for K_m can be evaluated as a projection over the two sub-spaces S_1 and S_* . Such that the dispersion given by $\text{disp}(K_m - P_{S_m}^{K_m})$ is minimized so that;

$$K_m^* = P_{S_m}^{K_m} = P_{S_1}^{K_m} + P_{S_*}^{K_m} = K_m + P_{S_m}^{K_m} \quad (4)$$

When n is assumed to be finite large data, so that the coefficients $\{a_v: v \geq m + 1\}$ are estimated such that the dispersion error of the estimate is minimized. This is achieved as follows:

We use equations (2) and (3) above to estimate the dispersion, such that the

$$\text{disp} \{K_m - P_{S_m}^{K_m}\} \text{ is minimized i.e. } K_m^* = P_{S_m}^{K_m} = P_{S_1}^{K_m} + P_{S_*}^{K_m} = K_m + P_{S_m}^{K_m} \quad (5)$$

$$\text{But } P_{S_m}^{K_m} = \left\{ \sum_{v=m+1}^n \xi_v (K_v - K_v); \text{disp} (K_m - P_{S_m}^{K_m}) \right\} \quad (6)$$

Squaring both sides and taking the expectations, we obtain the dispersion error as;

$$\text{disp } X_m = E(K_m - K_m^*) = \left\{ (K_m - \hat{K}_m) - \sum_{v=m+1}^n \xi_v (K_v - \hat{K}_m) \right\}^2 \quad (7)$$

By minimizing the dispersion with respect to the coefficients (differentiating with respect to ξ_v and solving for ξ_v), we should obtain the coefficients ξ_v , for $v \geq m + 1$, which are used in estimating the missing values. The missing value at point k_v is estimated as;

$$\hat{K}_m^* = \hat{k}_m + \sum_{v=m+1}^n \xi_v (k_v - \hat{k}_v) \quad (8)$$

2. ESTIMATION OF THE IMPUTATION ESTIMATORS

2.1 Derivation of ARMA Process Imputation Estimators

Lemma

Suppose k_t is non-determined stationary process defined on the probability space (Ω, β, ρ) . Then the subspace Q_m^* is the direct sum of subspaces Q_1 and Q_* .

Theorem 1

The imputation optimal interpolation estimator for ARMA (1,1) process is given by;

$$\hat{k}_m = \hat{\theta}_1 k_{t-1} + \hat{\varphi}_1 e_{t-1} + \sum_{v=m+1}^n (\theta_1 + \varphi_1) - \frac{(\theta_1 + \varphi_1)^{2(v-m)}}{2} \left(\frac{n}{n-2} \right)_{v-1} (k_v - \hat{k}_v) \quad (1)$$

Proof

The stationary ARMA (1,1) process is given by,

$$k_t = \theta_1 k_{t-1} + \varphi_1 \varepsilon_{t-1} + \varepsilon_t \text{ where } \varepsilon_t \sim t(0,1) \quad (2)$$

Getting the recursive form of equation (9) above, we obtain

$$k_m = \sum_{i=1}^{\infty} \left\{ \prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{t-1} + \varepsilon_t \right\} \quad (3)$$

Obtaining the r-step future predicted form of the above equation (12), we write it as,

$$k_{m+r} = \sum_{i=1}^{\infty} \left\{ \prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{t+r-1} + \varepsilon_{t+r} \right\} \quad (4)$$

Or if we set $v = r + h$, then we can rewrite the above equation (4) as,

$$k_v = \sum_{i=1}^{\infty} \left\{ \prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{v-1} + \varepsilon_v \right\} \quad (5)$$

Obtaining r-step future predicted error of the above equation (v), we obtain,

$$k_v - \hat{k}_v = \sum_{i=1}^{r-1} \left\{ \prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{v-1} + \varepsilon_v \right\} \quad (6)$$

From the above statement we know that the dispersion is given as,

$$\text{disp } k_m = E(k_m - \hat{k}_m)^2 = E\left\{k_m - \hat{k}_m - \sum_{v=m+1}^n \xi_v(k_v - \hat{k}_v)\right\}^2 \quad (7)$$

If we simplify the second part of the equation (16) above, then we obtain

$$\text{disp } k_m = E(k_m - \hat{k}_m)^2 - 2E\left\{\sum_{v=m+1}^n \xi_v(k_m - \hat{k}_m)(k_v - \hat{k}_v)\right\} + \left\{\sum_{v=m+1}^n \xi_v(k_v - \hat{k}_v)\right\}^2 \quad (8)$$

Substituting the above equation (15) into equation (17) we obtain the dispersion as;

$$\begin{aligned} \text{disp } k_m = E(k_m - \hat{k}_m)^2 - 2E\left\{(\mathcal{E}_m) \left(\sum_{i=1}^{r-1} \left\{\prod_{j=1}^i (\theta_1 + \varphi_1) \mathcal{E}_{v-1} + \mathcal{E}_v\right\}\right)\right\} \\ + E\left\{\sum_{v=m+1}^n \xi_v \left(\sum_{i=1}^{r-1} \left\{\prod_{j=1}^i (\theta_1 + \varphi_1) \mathcal{E}_{v-1} + \mathcal{E}_v\right\}\right)\right\}^2 \end{aligned} \quad (9)$$

Evaluating the above equation (18) then we can have;
The first term to be,

$$k_m = E(k_m - \hat{k}_m)^2 = E(\mathcal{E}_m)^2$$

The second term given by,

$$-2E\left[(\mathcal{E}_m) \sum_{v=m+1}^n \xi_v \left(\sum_{i=1}^{r-1} \left\{\prod_{j=1}^i (\theta_1 + \varphi_1) \mathcal{E}_{v-1} + \mathcal{E}_v\right\}\right)\right]$$

Which can be evaluated further to obtain,

$$-2E\left[\begin{aligned} &\xi_{m+1}(\theta_1 + \varphi_1)\mathcal{E}_m \cdot \mathcal{E}_m + \xi_{m+1}\mathcal{E}_{m+1} \\ &+ \xi_{m+2}(\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1)\mathcal{E}_{m+1} \cdot \mathcal{E}_m + \xi_{m+2}\mathcal{E}_{m+2} \\ &+ \xi_{m+3}(\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1)\mathcal{E}_{m+2} \cdot \mathcal{E}_m \\ &+ \xi_{m+3}\mathcal{E}_{m+3} + \dots \end{aligned}\right]$$

Or

$$-2E\left[\begin{aligned} &\xi_{m+1}(\theta_1 + \varphi_1)\mathcal{E}_m^2 + \xi_{m+1}\mathcal{E}_{m+1} \\ &+ \xi_{m+2}(\theta_1 + \varphi_1)^2\mathcal{E}_{m+1} \cdot \mathcal{E}_m + \xi_{m+2}\mathcal{E}_{m+2} \\ &+ \xi_{m+3}(\theta_1 + \varphi_1)^3 \cdot \mathcal{E}_{m+2} \cdot \mathcal{E}_m + \xi_{m+3}\mathcal{E}_{m+3} + \dots \end{aligned}\right]$$

Which can be simplified further to be,

$$-2E\{\xi_{m+1}(\theta_1 + \varphi_1)\mathcal{E}_m^2\}$$

The third term given by,

$$+E \left[\sum_{v=m+1}^n \xi_v \left(\sum_{i=1}^{r-1} \left\{ \prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{v-1} + \varepsilon_v \right\} \right) \right]^2$$

Can be evaluated to be,

$$+E \left[\begin{array}{c} \xi_{m+1}(\theta_1 + \varphi_1) \varepsilon_m \cdot \varepsilon_m + \xi_{m+1} \varepsilon_{m+1} \\ + \xi_{m+2}(\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1) \varepsilon_{m+1} \cdot \varepsilon_m + \xi_{m+2} \varepsilon_{m+2} \\ + \xi_{m+3}(\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1) \cdot (\theta_1 + \varphi_1) \varepsilon_{m+2} \cdot \varepsilon_m \\ + \xi_{m+3} \varepsilon_{m+3} + \dots \end{array} \right]^2$$

Or

$$+E \left\{ \begin{array}{c} \xi_{m+1}^2(\theta_1 + \varphi_1)^2 \varepsilon_m^2 \varepsilon_m^2 + \xi_{m+1}^2 \varepsilon_{m+1}^2 \\ + \xi_{m+2}^2(\theta_1 + \varphi_1)^4 \varepsilon_{m+1}^2 \varepsilon_m^2 + \xi_{m+2}^2 \varepsilon_{m+2}^2 \\ + \xi_{m+3}^2(\theta_1 + \varphi_1)^6 \varepsilon_{m+2}^2 \varepsilon_m^2 + \xi_{m+3}^2 \varepsilon_{m+3}^2 + \dots \end{array} \right\}$$

We can simplify the above equation of the third term to obtain,

$$+E \left[\varepsilon_m^2 \sum_{v=m+1}^n \xi_v (\theta_1 + \varphi_1)^{2(v-m)} \varepsilon_{v-1}^2 + \sum_{v=m+1}^n \xi_v^2 \varepsilon_v^2 \right]$$

We put all terms of the above equation (15) to obtain the whole equation under the dispersion as;

$$k_m = \left[\begin{array}{c} (\varepsilon_m)^2 - 2E\{\xi_{m+1}(\theta_1 + \varphi_1)\varepsilon_m^2\} \\ + \left\{ \varepsilon_m^2 \sum_{v=m+1}^n \xi_v (\theta_1 + \varphi_1)^{2(v-m)} \varepsilon_{v-1}^2 + \sum_{v=m+1}^n \xi_v^2 \varepsilon_v^2 \right\} \end{array} \right] \quad (10)$$

Substituting for the errors in the above equation (19) using the assumption of the distribution taken, we can have the dispersion as;

$$\text{disp}(\hat{x}_m) = \left(\left[\begin{array}{c} \frac{n}{n-2} - 2 \left\{ \xi_{m+1}(\theta_1 + \varphi_1) \frac{n}{n-2} \right\} + \frac{n}{n-2} \\ \sum_{v=m+1}^n \xi_v (\theta_1 + \varphi_1)^{2(v-m)} \left(\frac{n}{n-2} \right)_{v-1} + \sum_{v=m+1}^n \xi_v^2 \left(\frac{n}{n-2} \right)_v \end{array} \right] \right) \quad (11)$$

Differentiating the above equation (11), then;

$$\frac{\partial}{\partial \xi_v} \{\text{disp} x_m\} = \left(\begin{array}{c} \left(\frac{\partial}{\partial \xi_v} \left\{ \frac{n}{n-2} \right\} - 2 \frac{\partial}{\partial \xi_v} \left\{ \xi_{m+1}(\theta_1 + \varphi_1) \frac{n}{n-2} \right\} \right) \\ \left(+ \frac{\partial}{\partial \xi_v} \left\{ \frac{n}{n-2} \sum_{v=m+1}^n \xi_v (\theta_1 + \varphi_1)^{2(v-m)} \left(\frac{n}{n-2} \right)_{v-1} \right\} + \frac{\partial}{\partial \xi_v} \left\{ \sum_{v=m+1}^n \xi_v^2 \left(\frac{n}{n-2} \right)_v \right\} \right) \end{array} \right) \quad (12)$$

Evaluating the above equation (12) under the derivative and equating it to zero we obtain;

$$-2 \left\{ (\theta_1 + \varphi_1) \frac{n}{n-2} \right\} + \frac{n}{n-2} (\theta_1 + \varphi_1)^{2(v-m)} \left(\frac{n}{n-2} \right)_{v-1} + 2\xi \left(\frac{n}{n-2} \right) = 0 \quad (13)$$

Making ξ_v to be the subject of the above equation (13), then we obtain;

$$\xi_v = (\theta_1 + \varphi_1) - \frac{(\theta_1 + \varphi_1)^{2(v-m)}}{2} \left(\frac{n}{n-2} \right)_{v-1} \quad (14)$$

Substituting the above equation (14) on the expression of ξ_v into the interpolation equation (8), we get the optimal interpolator to be;

$$\hat{k}_m = \hat{\theta}_1 k_{t-1} + \hat{\varphi}_1 e_{t-1} + \sum_{v=m+1}^n (\theta_1 + \varphi_1) - \frac{(\theta_1 + \varphi_1)^{2(v-m)}}{2} \left(\frac{n}{n-2} \right)_{v-1} (k_v - \hat{k}_v) \quad (28)$$

Theorem 2

The imputation optimal interpolation estimator for ARMA (2, 1) process is given by

$$\hat{k}_t = \theta_2 \hat{k}_{t-2} + \theta_1 \hat{k}_{t-1} + \varphi_1 \varepsilon_{t-1} + \sum_{v=m+1}^n \frac{\theta_2}{(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1} (k_v - \hat{k}_v) \quad (1)$$

Proof

A stationery ARMA (2,1) process is given as

$$k_t = \theta_2 k_{t-2} + \theta_1 k_{t-1} + \varphi_1 \varepsilon_{t-1} + \varepsilon_t \quad \text{Where } \varepsilon_t \sim t(0,1) \quad (2)$$

We can write the recursive form of the above equation (2) to be;

$$k_t = \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_2) \varepsilon_{t-2j} \right] + \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{t-j} \right] + \varepsilon_t \quad (3)$$

The r-step forecast of the above process in the equation (3) is given by;

$$k_{t+r} = \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_2) \varepsilon_{t+r-2j} \right] + \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{t+r-j} \right] + \varepsilon_{t+r} \quad (4)$$

Or if we set $v = t + r$, then we can rewrite the above equation (3) to be,

$$k_v = \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_2) \varepsilon_{v-2j} \right] + \sum_{i=1}^{\infty} \left[\prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{v-j} \right] + \varepsilon_v \quad (5)$$

The future prediction error of equation (5) is given as,

$$k_v - \hat{k}_v = \sum_{i=1}^{v-1} \left[\prod_{j=1}^i (\theta_2) \varepsilon_{v-2j} \right] + \sum_{i=1}^{v-1} \left[\prod_{j=1}^i (\theta_1 + \varphi_1) \varepsilon_{v-j} \right] + \varepsilon_v$$

(6)

From the above lemma, we know that the dispersion is given by,

$$\text{disp } k_m = E(k_m - \hat{k}_m)^2 = E\{k_m - \hat{k}_m - \sum_{v=m+1}^n \xi_v (k_v - \hat{k}_v)\}^2 \quad (7)$$

If we simplify the second part of the equation (7) above, then we obtain the following expression

$$\text{disp } k_m = E(k_m - \hat{k}_m)^2 - 2 \left\{ \sum_{v=m+1}^n \xi_v (k_m - \hat{k}_m) (k_v - \hat{k}_v) \right\} + E \left\{ \sum_{v=m+1}^n \xi_v (k_v - \hat{k}_v) \right\}^2 \quad (8)$$

Substituting the above equation (6) in to the dispersion given by equation (8), we obtain the dispersion given by $\text{disp } k_m$ to be,

$$\begin{aligned} \text{disp } k_m = E(k_m - \hat{k}_m)^2 - 2E(e_m) & \left[\sum_{v=m+1}^n \xi_v (e_m) \left(\frac{\sum_{i=1}^{\infty} [\prod_{j=1}^i (\theta_2)^{\varepsilon_{v-2j}}]}{+\sum_{i=1}^{\infty} [\prod_{j=1}^i (\theta_1 + \varphi_1)^{\varepsilon_{v-j}}] + \varepsilon_v} \right) \right] \\ + E & \left[\sum_{v=m+1}^n \xi_v \left(\frac{\sum_{i=1}^{\infty} [\prod_{j=1}^i (\theta_2)^{\varepsilon_{v-2j}}]}{+\sum_{i=1}^{\infty} [\prod_{j=1}^i (\theta_1 + \varphi_1)^{\varepsilon_{v-j}}] + \varepsilon_v} \right) \right]^2 \end{aligned} \quad (9)$$

If we evaluate further the above equation (9), we obtain that

The first term

$$E(k_m - \hat{k}_m)^2 = E(\varepsilon_m)^2$$

The second term is,

$$-2E[e_m] \left[\begin{array}{l} \xi_{m+1} \theta_2 \varepsilon_{m-1} + \xi_{m+1} (\theta_1 + \varphi_1) \varepsilon_{m-1} + \xi_{m+1} \varepsilon_{m-1} \\ \xi_{m+2} \theta_2 \varepsilon_m + \xi_{m+2} (\theta_1 + \varphi_1)^2 \varepsilon_{m+1} + \xi_{m+2} \varepsilon_{m+2} \\ \xi_{m+3} \theta_3^2 \theta_2 \varepsilon_{m+1} + \xi_{m+3} (\theta_1 + \varphi_1)^2 \varepsilon_{m+2} + \xi_{m+3} \varepsilon_{m+3} + \dots \end{array} \right]$$

Which can be simplified further to be,

$$-2E \left[\sum_{v=m+2}^n (\varepsilon_m^2) \xi_v \theta_2 \right]$$

The third term can be expressed as,

$$+ \sum_{v=m+1}^n \xi_v^2 \cdot \sum_{i=1}^{\infty} \left[\prod_j^i (\theta_2)^2 \varepsilon_{v-2j}^2 \right] + \sum_{v=m+1}^n \xi_v^2 \cdot \sum_j^n \left[\prod_i^j (\theta_1 + \varphi_1)^2 \varepsilon_{v-i}^2 \right] + \sum_{v=m+1}^n \xi_v^2 \varepsilon_v^2$$

Can be evaluated as,

$$+ E \left[\begin{array}{l} \xi_{m+1}^2 \cdot (\theta_2)^2 \varepsilon_{m-1}^2 + \xi_{m+1}^2 \cdot (\theta_1 + \varphi_1)^2 \varepsilon_m^2 + \xi_{m+1}^2 \varepsilon_{m+1}^2 \\ \xi_{m+2}^2 \cdot (\theta_2)^4 \varepsilon_m^2 + \xi_{m+2}^2 \cdot (\theta_1 + \varphi_1)^4 \varepsilon_{m+1}^2 + \xi_{m+2}^2 \varepsilon_{m+2}^2 \\ \xi_{m+3}^2 \cdot (\theta_2)^8 \varepsilon_{m+1}^2 + \xi_{m+3}^2 \cdot (\theta_1 + \varphi_1)^6 \varepsilon_{m+1}^2 + \xi_{m+3}^2 \varepsilon_{m+3}^2 + \dots \end{array} \right]$$

Which can be written as,

$$+E\left(\sum_{v=m+1}^n \xi_v^2 \cdot (\theta_2)^{2(v-m)} \varepsilon_v^2 + \sum_{v=m+1}^n \xi_v^2 \cdot (\theta_1 + \varphi_1)^{2(v-m)} \varepsilon_v^2 + \sum_{v=m+1}^n \xi_v^2 \varepsilon_v^2 + \dots\right)$$

Which can be generalized as;

$$+E\left\{\sum_{v=m+1}^{v-m} \xi_v^2 \varepsilon_v^2 [(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1]\right\}$$

Putting all terms of equation (9) together, then we can have,

$$\text{disp } k_t = E(\varepsilon_m)^2 - 2E\left[\sum_{v=m+2}^n (\varepsilon_m^2) \xi_v \theta_2\right] + E\left\{\sum_{v=m+1}^n \xi_v^2 \varepsilon_v^2 [(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1]\right\} \quad (10)$$

Substituting for the above equation (10) for the errors using the characteristics of the assumed distribution, then we will have the dispersion as,

$$\text{disp } k_t = \frac{n}{n-2} - 2\frac{n}{n-2}\left[\sum_{v=m+2}^n \xi_v \theta_2\right] + \frac{n}{n-2}\left\{\sum_{v=m+1}^n \xi_v^2 [(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1]\right\} \quad (11)$$

Differentiating the above equation with respect to ξ_v

$$\begin{aligned} \frac{\partial}{\partial \xi_v} (\text{disp } k_t) &= \frac{\partial}{\partial \xi_v} \left(\frac{n}{n-2}\right) - 2\frac{\partial}{\partial \xi_v} \left[\frac{n}{n-2} \left[\sum_{v=m+2}^n \xi_v \theta_2\right]\right] \\ &+ \frac{\partial}{\partial \xi_v} \left[\frac{n}{n-2} \left\{\sum_{v=m+1}^n \xi_v^2 [(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1]\right\}\right] \end{aligned} \quad (12)$$

Setting it to zero, we obtain;

$$\xi_v = \sum_{v=m+1}^n \frac{\theta_2}{(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1} \quad (13)$$

Substituting the above equation (13) on the expression of ξ_v into the interpolation equation (7), we get the optimal interpolator to be;

$$\hat{k}_t = \theta_2 \hat{k}_{t-2} + \theta_1 \hat{k}_{t-1} + \varphi_1 \varepsilon_{t-1} + \sum_{v=m+1}^n \frac{\theta_2}{(\theta_2)^{2(v-m)} + (\theta_1 + \varphi_1)^{2(v-m)} + 1} (k_v - \hat{k}_v)$$

3. RESULTS AND DISCUSSION

Synthetic Data and Simulation

We generated 1000 samples through simulation using R statistical package version 4.42. The simulation of samples was done taking into consideration the student's t assumption. We created 100 missing values within the simulated data sets. The missing mechanism for the missing data in the simulated data was missing at random (MAR) mechanism. The codes for simulations are provided here in.

The distribution of the generated data is given below.

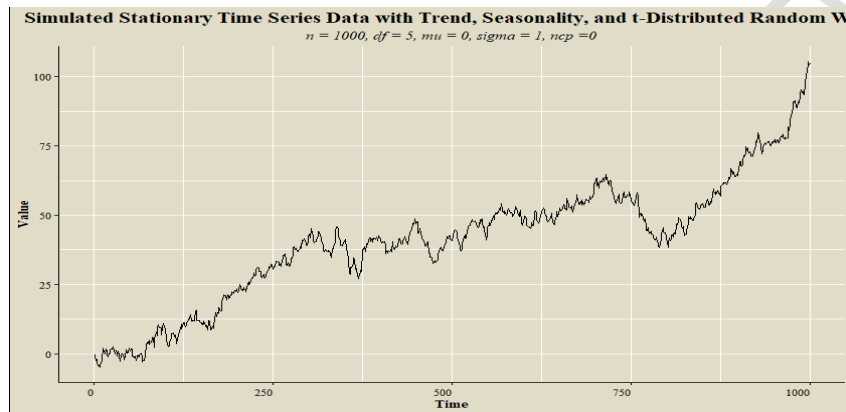


Figure 1.

Introducing 100 random missing data into the simulated data, we get an illustration like the one below.

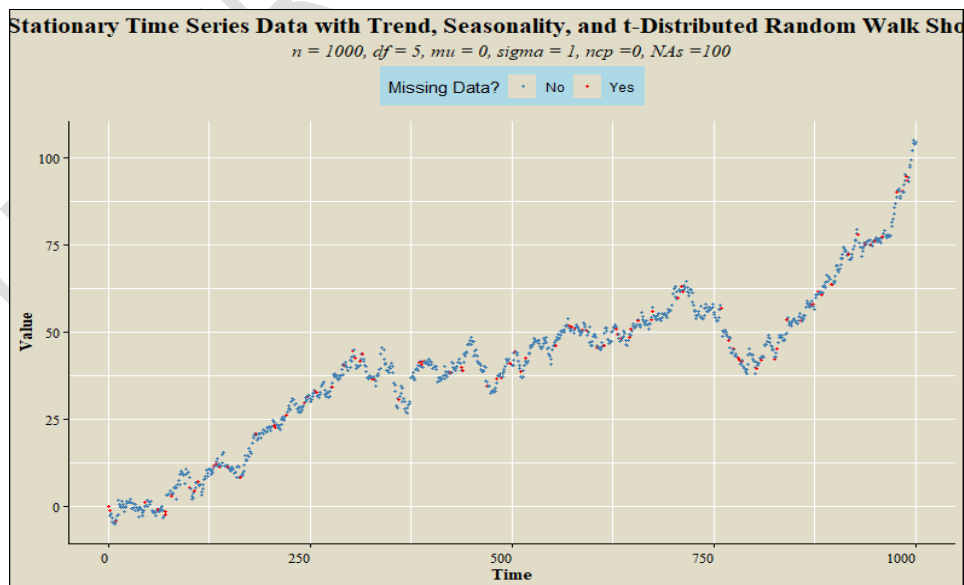


Figure 2.

Imputing the created missing values using the above model ARMA (1,1), we then observe that the imputation is done well with the derived estimating function.

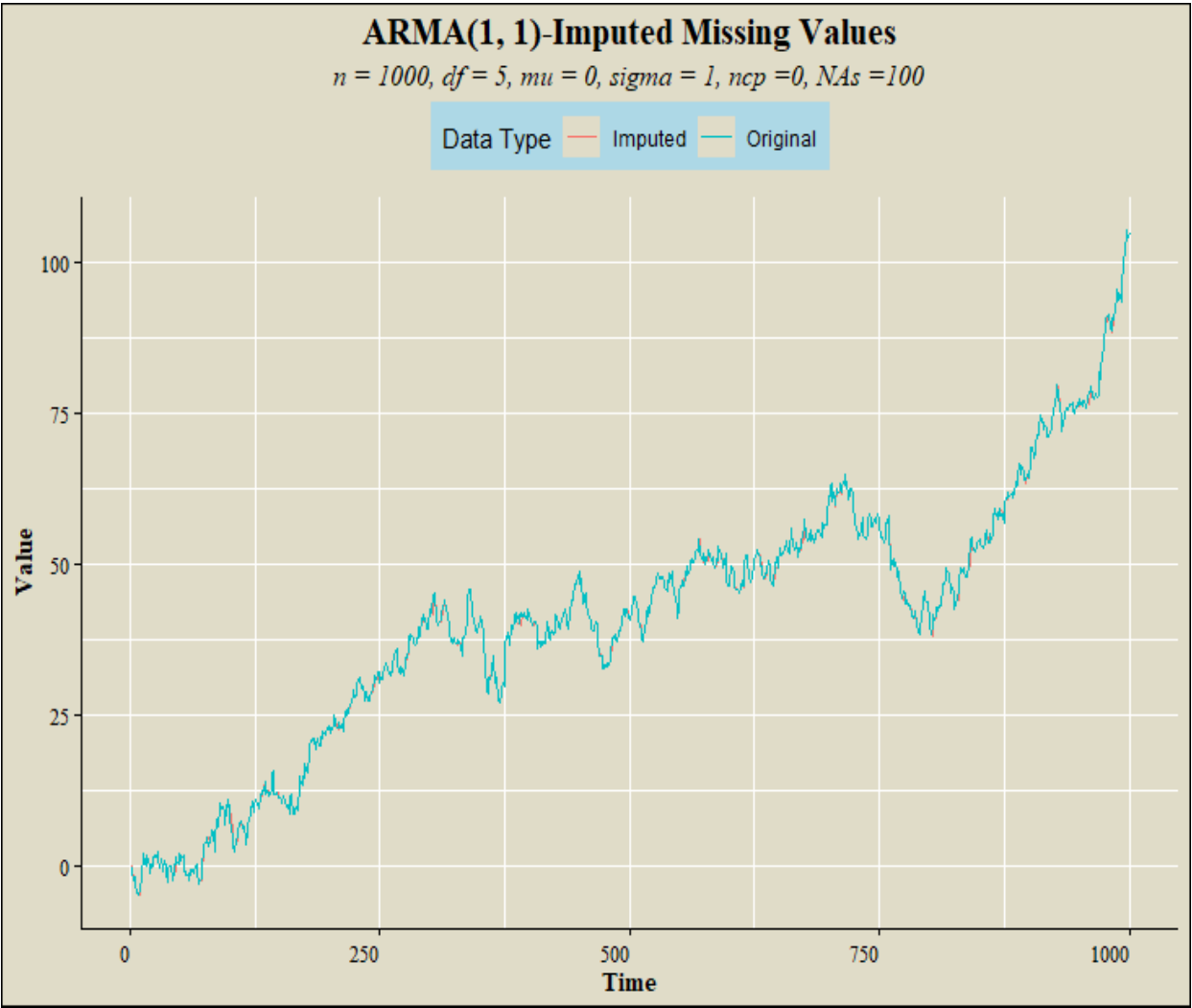


Figure 3.

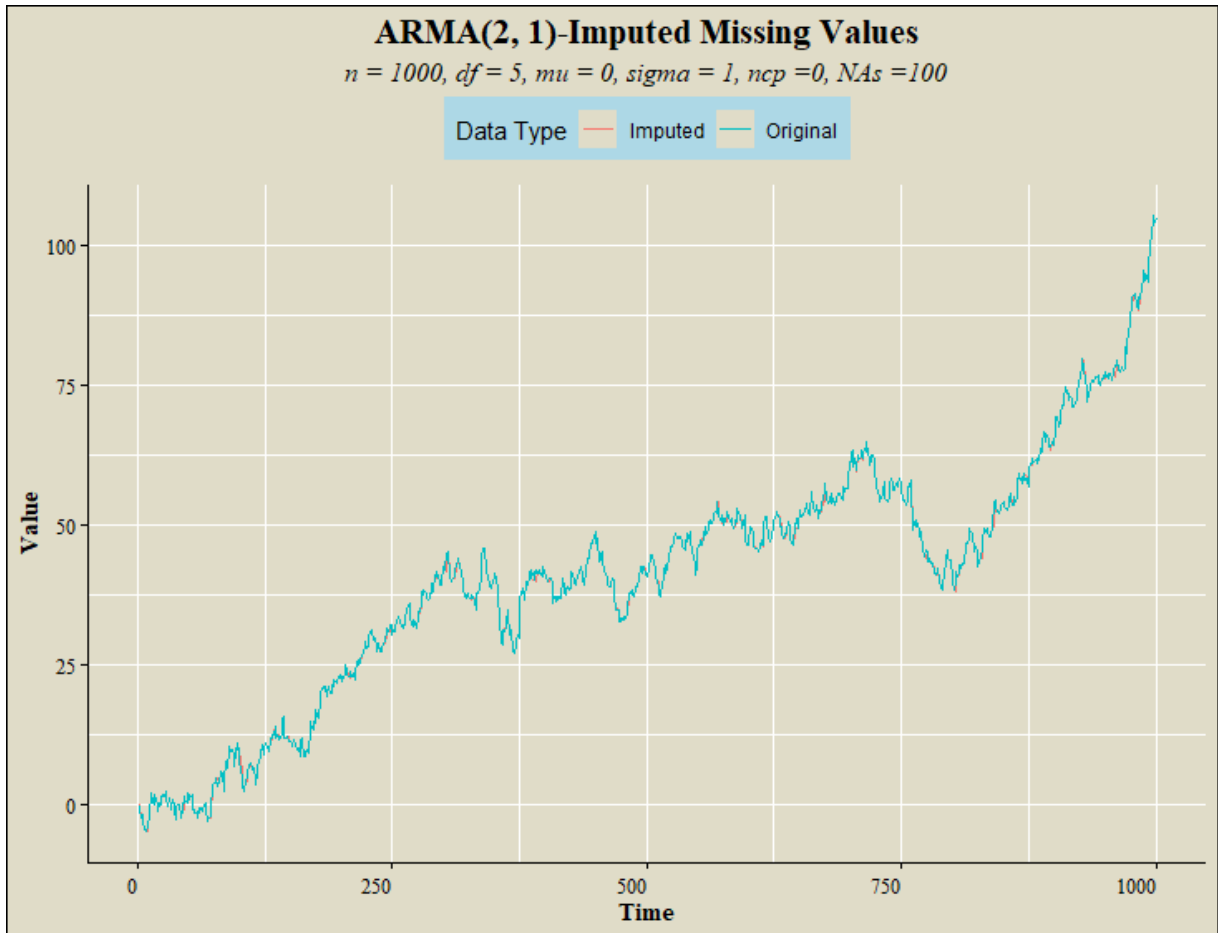


Figure 4.

The imputation metrics were also calculated for other imputation techniques, besides the ones introduced in the above. We see that

4. CONCLUSION

The derived and adjusted ARMA imputation estimators, were able to compete favorably with the state of the art imputation techniques like ANN, a machine learning technique that has taken the center stage at the prediction and forecasting, Kalman filters and the KNN.

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APPENDIX

```
# load packages
library(ggplot2)
library(forecast)
# set seed for reproducibility
set.seed(44)
# set parameters
n <- 1000 # number of records
df <- 5 # degrees of freedom the t-distribution
mu <- 0 # mean of the time series
sigma <- 1 # standard deviation of the time series
ncp <- 0 # non-centrality parameter (optional)
NAs <- 100 # number of missing records
# create time indices
time_index <- 1:n
# generate trend and seasonal components
trend <- 0.1 * time_index
seasonal_component <- sin(2 * pi * time_index / 12)
# simulate the random walk component using t-distribution
random_walk <- cumsum(rt(n, df) * sigma + mu)
# combine the components and create a time series
ts_data <- ts(trend + seasonal_component + random_walk, start = c(2000, 1), frequency =
12)
# Convert time series to data frame
df_ts_data <- data.frame(Time = time_index, Value = as.numeric(ts_data))
# custom graph theme
```

```

cust_theme <- function(){
  theme(plot.title = element_text(face = "bold",
  hjust = 0.5,
  size = 16,
  family = "serif",
  color = "black"),
  plot.subtitle=element_text(face = 'italic',
  hjust = 0.5,
  size = 12,
  family = "serif",
  color = 'black'),
  axis.title = element_text(face = "bold",
  size = 11.5,
  family = "serif",
  color = "black"),
  axis.text = element_text(face = "plain",
  size = 10,
  family = "serif",
  color = "black"),
  strip.text.x = element_text(face = "bold",
  size = 13.5,
  family = "serif",
  color = "black"),
  axis.text.x = element_text(angle = 0,
  hjust = 1,
  vjust = 0.5),
  plot.background = element_rect(fill = "#E0DCC8",
  color = "black",
  linewidth = 1),
  panel.background = element_rect(fill = "#E0DCC8"),
  axis.line = element_line(color = "black"),
  axis.ticks = element_line(color = "black"),

  legend.position = "top",
  legend.direction = "horizontal",
  legend.background = element_rect(fill = "lightblue")
  )
}
# visualize the time series using ggplot2
line_graph_t_series <- ggplot(df_ts_data, aes(x = Time, y = Value)) +
  geom_line() +
  labs(title = "Simulated Non-Stationary Time Series Data with Trend, Seasonality
  , and t-Distributed Random Walk",
  x = "Time",
  y = "Value",
  subtitle = paste0("n = ", n,
  ", df = ", df,
  ", mu = ", mu,
  ", sigma = ", sigma,
  ", ncp =", ncp)) +
  cust_theme()
# print visual
line_graph_t_series

```

```

# introduce missing data at random (data frame), ensuring reproducibility
set.seed(44)
missing_indices <- sample(1:n, NAs)
df_ts_data$Value_miss <- df_ts_data$Value
df_ts_data$Value_miss[missing_indices] <- NA
# introduce missing data at random (non-data frame), ensuring reproducibility
ts_data_miss <- ts_data
ts_data_miss[missing_indices] <- NA
# visualize missing Values
df_ts_data <- df_ts_data |>
dplyr::mutate(Missing = dplyr::if_else(!is.na(Value_miss), "No",
dplyr::if_else(is.na(Value_miss), "Yes", "")))
(line_graph_t_series_miss <- df_ts_data |>
ggplot(aes(x = Time,
y = Value,
color = Missing)) +
geom_jitter(size = 1, pch = 20) +
scale_color_manual(values = c("Yes" = "red", "No" = "steelblue")) +
labs(title = "Simulated Non-Stationary Time Series Data with Trend, Seasonality,
and t-Distributed Random Walk Showing Missing Values",
x = "Time",
y = "Value",
color = "Missing Data?",
subtitle = paste0("n = ", n,
", df = ", df,
", mu = ", mu,
", sigma = ", sigma,
", ncp = ", ncp,
", NAs = ", NAs)) +
cust_theme())
# save data into disc
data_dir <- paste0(getwd(), "/SimulatedData")
if(!dir.exists(data_dir)){
dir.create(data_dir)
} else {print("Directory Exists!")}
write.csv(x = df_ts_data,
file = paste0(data_dir, "/simulated_t_series_data.csv"),
row.names = FALSE)

```