

## Original Research Article

# APPLICATION OF THE BOX AND JENKINS METHOD TO THE TIME SERIES OF SUCCESS RATES OF STATE EXAMINATIONS IN THE DEMOCRATIC REPUBLIC OF CONGO FROM 1967 TO 2021

### ABSTRACT

In this article, we present an analysis conducted on the time series of success rates in State Examinations in the Democratic Republic of Congo (DRC) since its inception in 1967 until 2021, following the BOX and JENKINS methodology, which allows for obtaining a parsimonious ARIMA family model capable of representing the series.

The best ARIMA model obtained from the studied series is ARIMA (0, 1, 2); which proved to be the only one to respect all the steps of the Box-Jenkins method for the series of success rates in the State Examinations in the DRC. We used Eviews 12.0 software and applied first differentiating. Our forecast results show upward trends and almost stability from 2023 to 2025.

**Keywords :** Time series, Box-Jenkins methodology, Success rates, ARIMA models, Parsimonious.

### 1. INTRODUCTION

Currently throughout the Universe, each State is organizing itself to put in place, according to its needs, the type of education desired. In the Democratic Republic of Congo, we inherited the system from our colonizer, which is none other than Belgium. The time has come for us to give our voice to the end of secondary education evaluation system in the DRC, which is the one and only way to obtain the state diploma.

The importance of education no longer needs to be demonstrated. We emphasize the aspect that it is the basis of development and without it, the individual remains in the wild.

We have chosen for our part, an analysis based on the success rates in the so-called state exams, which we have collected since its establishment in 1967 until 2021. In total 55 observations which, taken together, form a series simple chronology.

We used the Box and Jenkins methodology. This justifies the originality of this study and predicting these rates in the short term over a period of four years.

The objective pursued is to analyze this chronicle and above all to predict as said above.

The approach pursued is that which consisted of collecting at the General Inspection of Primary, Secondary and Technical Education (EPST), the different success rates

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in state exams which form a chronicle whose size is 55 observations and which then was analyzed following the methodology of BOX and JENKINS.

The tool used is the Eviews 12 software with the primary differentiation technique

The forecasts obtained are very important for a view on the future of success rates in the State Examination which is the one and only instrument for evaluating National Education at the end of the secondary cycle in the DRC.

The structure of this article includes the following points:

- ✓ Introduction,
- ✓ Presentation of the Box and Jenkins method,
- ✓ Data presentation,
- ✓ Data analysis and interpretation,
- ✓ Prediction,

A conclusion ends the article

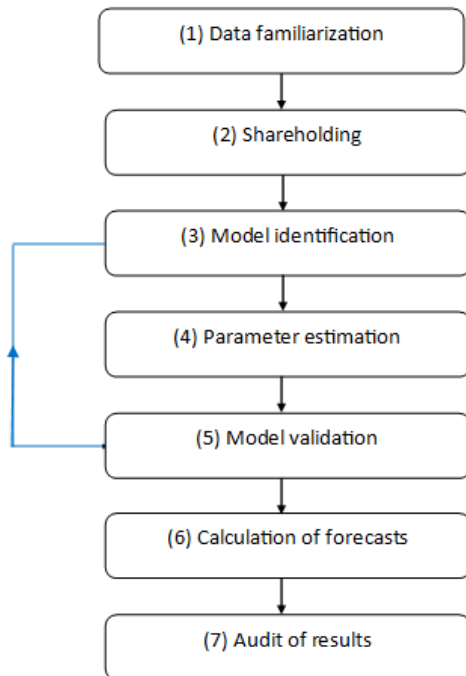
## 2. PRESENTATION OF THE BOX-JENKINS METHOD

Long before the appearance the Box-Jenkins method, the type of forecast most often known in statistical practice, made it possible to accept the existence of a fundamental law independent of the series, represented by these data with a random character; it consisted of isolating the statistical characteristics of the series in order to use it to project the future [5].

In fact, statistical analysis boiled down to the attempt to adjust observations using a model established in advance; Thus, to overcome this problem, Professors George Box and Gwilym Jenkins proposed together in 1976 a methodological approach offering the opportunity to judge whether a law of variation is satisfactory to the extent possible ; it is an iterative approach which consists of identifying an adequate model capable of representing the phenomenon studied. These steps are represented by the following flowchart. [2]

Flowchart 1: Flow chart showing Box-Jenkins Method

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Comments to make on this flowchart are as follows:

(1) At the data familiarization stage, this involves:

- ✓ Represent data graphically;
- ✓ Check if there is any missing data;
- ✓ Check for outliers;
- ✓ Identify the trend;
- ✓ Check stationarity;

(2) Stationarization: stationarize the series using difference or using logarithmic transformation until obtaining the stationarized series.

(3) identification of the model: here consists of determining p and q which correspond to the series to have the number of variables at each polynomial. We used the parsimony approach, i.e. taking p and q as small as possible so as not to lose a lot of information.

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(4) Parameter estimation: The best method must be used. We used the ordinary least squares method therefore:

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$$\rho_k = \frac{\text{cov}(X_t, X_{t-1})}{\text{var}(X_t)} = \theta_1^k \frac{\text{var}(X_t)}{\text{var}(X_t)} = \theta_1^k \quad (1)$$

It should be noted that to verify the stationarity of the AR(1) process, it is necessary and sufficient that  $|\theta_1| < 1$  which corresponds to the exponential decay ( $0 < \theta_1 < 1$ )

Or ( $-1 < \theta_1 < 0$ ) oscillatory.

ProcessAR(p)

We either have a process (p), given by:

$$X_t = \theta_0 + \theta_1 X_{t-1} + \dots + \theta_p X_{t-p} + \varepsilon_t$$

With k=1,2, 3,...

For a process AR(1) ; we have

$$X_t = \theta_0 + \theta_1 X_{t-1} + \varepsilon_t$$

**-Theoretical simple autocorrelation function of a Process MA(1)**

Or the process MA(1):

$$X_t = \varphi_1 \varepsilon_{t-1} + \varepsilon_t \quad (2)$$

$$\rho_k = \begin{cases} \rho_0 = 1 \\ \rho_1 = \frac{\varphi_1}{1-\varphi_1^2} \\ \vdots \\ \rho_k = 0 \quad \forall k > 1 \end{cases} \quad (3)$$

We can see and generalize that for a process MA(q);

$$\rho_k \neq 0 \quad \text{For } k \leq q \text{ et } \rho_k = 0 \quad \text{for } k > q \quad (4)$$

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Identification is done by comparing the FAC empirical to the FAC theoretical

**- Partial autocorrelation function (PACF) for the parameter p**

The partial autocorrelation function is comparable to the simple autocorrelation function and also to the partial correlation between two variables  $X_t$  et  $X_{t-k}$  is this correlation considered after eliminating the effect of all intermediate lags i.e. the effect  $X_{t-1}, X_{t-2}, \dots, X_{t-k+1}$

From one FAC, we can find the FACP.

Here is the relationship;  $\theta_{11} = \rho_1$

(5)

$$\text{et } \theta_{22} = \frac{\rho_2 - \rho_1^2}{1 - \rho_1^2} \quad (6)$$

si la série est d'ordre k noté : AR(k), nous pouvons généralisée en :

$$\theta_{kk} = \frac{\rho_k - \sum_{i=1}^{k-1} \theta_{(k-1)i} \rho_{k-i}}{1 - \sum_{i=1}^{k-1} \theta_{(k-1)i} \rho_i} \quad \text{pour } k > 2 \quad (7)$$

$$\theta_{ki} = \theta_{(k-1)i} \theta_{(k-1)(k-i)} \quad (8)$$

The coefficients  $\theta_{kk}$  coming from the estimation of the models above are the values of the simple autocorrelation function of order  $k = 1, 2, 3, \dots$

If  $k = 1$

$$X_t = \theta_0 + \theta_{11} X_{t-1} + \varepsilon_t$$

If  $k = 2$ ; SO :

$$X_t = \theta_0 + \theta_{21} X_{t-1} + \theta_{22} X_{t-2} + \varepsilon_t$$

Si  $k = n$ ;

$$X_t = \theta_0 + \theta_{n1}X_{t-1} + \theta_{n2}X_{t-2} + \dots + \theta_{nn}X_{t-n} + \varepsilon_t \quad (9)$$

For a process AR(1)

$$X_t = \theta_0 + \theta_1 X_{t-1} + \varepsilon_t$$

With  $|\theta_1| < 1$

$$\theta_{kk} \begin{cases} \neq 0 \text{ pour } k = 1 \\ = 0 \forall k > 1 \end{cases} \quad (10)$$

For AR(p);

$$X_t = \theta_0 + \theta_1 X_{t-1} + \dots + \theta_p X_{t-p} + \varepsilon_t$$

$$\theta_{kk} \begin{cases} \neq 0 \text{ pour } k = 1, \dots, p \\ = 0 \forall k > p \end{cases} \quad (11)$$

NOTE :

The partial autocorrelation function (FACP) of a process MA(q) decays exponentially or oscillatorily. It is important to remember that the process MA(q) is always stationary. [2]

The study of autocorrelation functions allows us to identify the series in a family model ARMA, which is the adequate model, that is to say the model which has parameters significantly different from zero. The objective here is to determine the form of the model which models the series studied.

The involvement of a few tests will allow us to deseasonalize and stationarize the series to allow the application of the so-called Box-Jenkins methodology.

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(5) Model validation: We used several statistical tests as described in the text

(6) Calculates forecasts: To make forecasts, you must modify the size of the series, up to the forecast period envisaged. We should therefore use the forecast formula of an ARIMA model given by:

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$$DTREUSSIT. F_t = c + \sum_{i=1}^{p+d} \theta_i X_{t-i} + \varepsilon_t + \sum_{k=j}^q \varphi_j \varepsilon_{t-j}$$

And

$$\hat{X}_{n+h} = c + \sum_{i=1}^{p+d} \theta_i \hat{X}_{n+h-i} + \sum_{j=1}^q \varphi_j \hat{\varepsilon}_{n+h-j}$$

$$\text{If } k \geq h \text{ and } \hat{\varepsilon}_{n+h-k} = \begin{cases} \varepsilon_{n+h-k} \text{ si } k \geq h \\ 0 \text{ si } k < h \end{cases}$$

If we replace the data we have in the formula, we will have;

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For h=1

$$X_t = c + \sum_{k=1}^{0+1} \theta_k X_{t-k} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k}$$

$X_t = c + \sum_{k=1}^{0+1} \theta_k X_{t-k} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k}$  Which brings us to;

$$X_t = c + \theta_1 X_{t-1} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k} \text{ And}$$

$$\hat{X}_{55+1} = \sum_{k=1}^{0+1} \theta_k \hat{X}_{55+1-k} + c + \sum_{k=1}^2 \varphi_k \hat{\varepsilon}_{55+1-k}$$

$$\hat{X}_{56} = \theta_1 \hat{X}_{55} + c + \sum_{k=1}^2 \varphi_k \hat{\varepsilon}_{55-k}$$

With  $h \geq q$ ; we obtain a relation of the form;  $\hat{X}_{n+h} = c + \sum_{k=1}^{p+d} \theta_k \hat{X}_{n+h-k}$   
For  $h=2$ ,

$$\hat{X}_{57} = c + \sum_{k=1}^{0+1} \theta_k \hat{X}_{55+2-k}$$

$$\hat{X}_{57} = c + \theta_1 \hat{X}_{55+2-1}$$

$$\hat{X}_{57} = c + \theta_1 \hat{X}_{56}$$

Of this ; we can deduce;

For  $h=3$ ,

$$\hat{X}_{58} = c + \theta_1 \hat{X}_{57}$$

For  $h=4$ ,

$$\hat{X}_{59} = c + \theta_1 \hat{X}_{58}$$

### 3. PRESENTATION OF DATA

We were interested in collecting data on pass rates for state exams in the DRC, from its establishment in 1967 until 2021, which we present here in table no. 1 .

Table No. 1: Different success rates in State Exams from 1967 to 2021

( Source: Evaluation Office of the General Inspection of Primary, Secondary and Vocational Education of the DRC / Kinshasa)

YEAR	RATE	YEAR	RATE	YEAR	RATE	YEAR	RATE	YEAR	RATE
1967	70	1978	18	1989	51	2000	65	2011	69
1968	65	1979	42	1990	50	2001	33	2012	61
1969	56	1980	41	1991	67	2002	41	2013	47
1970	67	nineteen eighty	40	1992	46	2003	68	2014	54

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1971	58	1982	30	1993	38	2004	67	2015	60
1972	68	1983	31	1994	59	2005	66	2016	61
1973	53	1984	40	1995	64	2006	63	2017	65
1974	52	1985	46	1996	56	2007	38	2018	67
1975	64	1986	51	1997	66	2008	46	2019	70
1976	70	1987	46	1998	63	2009	62	2020	72
1977	51	1988	45	1999	63	2010	66	2021	58

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#### 4. APPLICATION OF THE BOX-JENKINS METHOD

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##### 4.1. Familiarization

The study of the series that we call "TREUSSIT" presupposes a study of the graph and the correlogram of the series which will make it possible to detect the existence of a trend and/or seasonality of the said series. We used the reviews12 software.

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##### 4.1.1 Graph and correlogram of the TREUSSIT series

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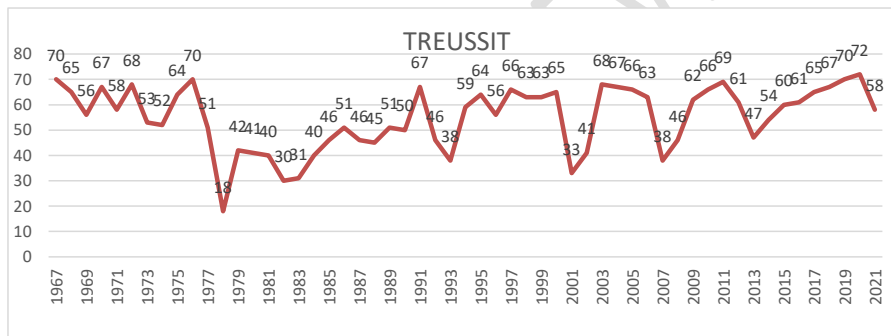


Figure 1: Graphical representation of the TREUSSIT series

Source: author using eviews 12 software

We see that there is not much fluctuation around the average. Therefore, the time series is not stationary because its trend is downward.

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Graphical examination does not always make it possible to determine with certainty the existence of a trend. In order to remove uncertainty, we use the correlogram and appropriate tests to prove or not the stationarity of the series.

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##### 4.1.2 Correlogram

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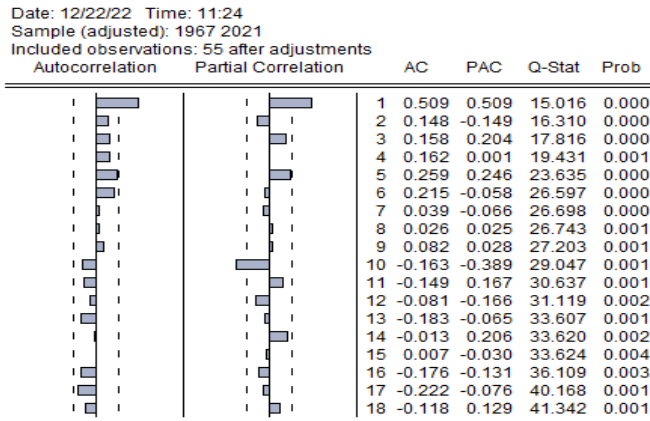


Figure 2: Correlogram of the TREUSSIT series

Source: Author using Eviews 12 software.

### Reading the correlogram

Columns AC and PAC are the results of the simple autocorrelation functions and the partial autocorrelation function which are represented respectively in the form of horizontal bars on the first two columns of this table.

Their confidence interval boundaries (dotted band) are stylized by vertical dotted lines. Each term that falls outside this interval is significantly different from 0 at the 5% threshold.

Q-Stat is the statistic of  $Q_{BP}$  lag k.

Prob is the critical probability of the test on the nullity of the coefficients  $\rho_k$

Finally the delay k is represented by the column between the PAC and AC columns

### Interpretation :

From the graphical representation we see the absence of many fluctuations around the trend. We can say that this diagram is that of a non-stationary series.

From the correlogram, we see that the probabilities are all less than 0.05 Hence the series is not stationary.

The series of State Exam pass rates is therefore non-stationary according to reading the graph and the correlogram of the raw series. It is up to us to prove this thesis or not by appropriate tests.

### 4.1.3 Study of the stationarity of the TREUSSIT series

There are several tests to check the stationarity of a series but the most used are that of Philippe Perron and the Dickey Fuller Augmenter Test (ADF). These tests not only help

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to test the stationarity of the series but also to show the best way to stationarize it in case of non-stationarity.

In this article, we use the Dickey Fuller Augmenter test, because according to reading the correlogram, we have probabilities all lower than 5%; it is therefore a process which does not follow the law of chance

We move on to hypothesis testing; we first pose the hypotheses then we apply the test at the 5% threshold.

**MODEL 3**

$$X_t = \alpha X_{t-1} + \beta t + c + \varepsilon_t$$

We make the following hypotheses:

- {  $H_0$ : absence of trend
- {  $H_1$ : presence of the trend

Null Hypothesis: TREUSSIT has a unit root  
 Exogenous: Constant, Linear Trend  
 Lag Length: 1 (Automatic - based on SIC, maxlag=10)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-4.497700	0.0037
Test critical values:		
1% level	-4.140858	
5% level	-3.496960	
10% level	-3.177579	

\*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation  
 Dependent Variable: D(TREUSSIT)  
 Method: Least Squares  
 Date: 01/03/23 Time: 13:56  
 Sample (adjusted): 1969 2021  
 Included observations: 53 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
TREUSSIT(-1)	-0.629654	0.139995	-4.497700	0.0000
D(TREUSSIT(-1))	0.170693	0.138336	1.233905	0.2231
C	29.51839	7.480724	3.945927	0.0003
@TREND("1967")	0.170432	0.099996	1.704386	0.0946
R-squared	0.305155	Mean dependent var		-0.132075
Adjusted R-squared	0.262613	S.D. dependent var		12.35070
S.E. of regression	10.60570	Akaike info criterion		7.633132
Sum squared resid	5511.565	Schwarz criterion		7.781834
Log likelihood	-198.2780	Hannan-Quinn criter.		7.690316
F-statistic	7.173099	Durbin-Watson stat		1.933723
Prob(F-statistic)	0.000437			

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Table 2: Model 3 of the ADF test on the TREUSSIT series

Source: Author using Eviews 12 software.

We have results found using the Eviews 12 software as follows:

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the probability linked to the trend;  $\text{prob}=0.0946 > 5\%$  therefore the trend is not significant and since the value of the ADF is greater than that of Mackinnon, that is to say  $-4.497700 < -3.496960$  at the threshold of 5%, we have the presence of the unit root, although the series does not contain a significant trend, it is not as stationary because of the presence of the unit root. From which we accept  $H_0$  which stipulates that there is the absence of the trend .

Now let's check the presence of the constant.

**MODEL 2**

$$X_t = \alpha X_{t-1} + c + \varepsilon_t$$

We make the following hypotheses:

- $\{ H_0: \text{absence of the constant}$
- $H_1: \text{presence of the constant}$

Null Hypothesis: TREUSSIT has a unit root  
 Exogenous: Constant  
 Lag Length: 0 (Automatic - based on SIC, maxlag=10)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-4.136688	0.0019
Test critical values:		
1% level	-3.557472	
5% level	-2.916566	
10% level	-2.596116	

\*Mackinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation  
 Dependent Variable: D(TREUSSIT)  
 Method: Least Squares  
 Date: 01/03/23 Time: 14:06  
 Sample (adjusted): 1968 2021  
 Included observations: 54 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
TREUSSIT(-1)	-0.481901	0.116494	-4.136688	0.0001
C	26.26446	6.567220	3.999326	0.0002
R-squared	0.247600	Mean dependent var		-0.222222
Adjusted R-squared	0.233131	S.D. dependent var		12.25156
S.E. of regression	10.72882	Akaike info criterion		7.620078
Sum squared resid	5985.592	Schwarz criterion		7.693744
Log likelihood	-203.7421	Hannan-Quinn criter.		7.648488
F-statistic	17.11219	Durbin-Watson stat		1.845171
Prob(F-statistic)	0.000129			

Table 3: Model 2 of the ADF test on the TREUSSIT series

Source: Author using Eviews 12 software

We have the probability linked to the constant given by:  $\text{prob}=0.0002$  which is less than 5%; the ADF equal to  $-4.13$  is lower than the Mackinnon critical value at the 5% threshold given by  $-2.91$  Hence; we accept  $H_1$  et we reject  $H_0$ . We maintain that there is the presence of a constant.

The model to be retained must necessarily have a constant.

Let's move on to checking the presence of the unit root.

**MODEL 1**

$$X_t = \alpha X_{t-1} + \varepsilon_t$$

- $\{ H_0: \text{presence of the unit root : the series is not stationary}$
- $H_1: \text{absence of the unit root : the series is stationary}$

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Null Hypothesis: TREUSSIT has a unit root  
 Exogenous: None  
 Lag Length: 4 (Automatic - based on SIC, maxlag=10)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-0.118200	0.6381
Test critical values:		
1% level	-2.612033	
5% level	-1.947520	
10% level	-1.612650	

\*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation  
 Dependent Variable: D(TREUSSIT)  
 Method: Least Squares  
 Date: 01/03/23 Time: 14:16  
 Sample (adjusted): 1972 2021  
 Included observations: 50 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
TREUSSIT(-1)	-0.003332	0.028193	-0.118200	0.9064
D(TREUSSIT(-1))	-0.313851	0.141883	-2.212045	0.0321
D(TREUSSIT(-2))	-0.585067	0.142515	-4.105294	0.0002
D(TREUSSIT(-3))	-0.303331	0.140699	-2.155894	0.0365
D(TREUSSIT(-4))	-0.342712	0.138753	-2.469950	0.0174
R-squared	0.304778	Mean dependent var		0.000000
Adjusted R-squared	0.242980	S.D. dependent var		12.49490
S.E. of regression	10.87142	Akaike info criterion		7.704791
Sum squared resid	5318.451	Schwarz criterion		7.895994
Log likelihood	-187.6198	Hannan-Quinn criter.		7.777602
Durbin-Watson stat	1.953591			

Table 4: Model 1 of the ADF test on the TREUSSIT series

Source: Author using eviews 12 software

The series is stationary because the ADF value is greater than the Mackinnon critical value. Which means,

$-0.118200 > -1.947520$ , then we accept  $H_0$  and reject  $H_1$ , so the series is not stationary, we are facing a DS model with drift.

We know that; To make a DS model process stationary, we will go through the differentiation method of order 1,2,... until the series becomes stationary. Here we apply the ADF test to the series in first differentiation.

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#### 4.1.4 TREUSSIT STATIONARIZATION

Hypotheses :

$H_0$ : The series is not stationary

$H_1$ : the series is stationary

The application of the ADF Test to the first differentiation series

Null Hypothesis: DTREUSSIT has a unit root  
 Exogenous: None  
 Lag Length: 3 (Automatic - based on SIC, maxlag=10)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-6.661759	0.0000
Test critical values:		
1% level	-2.612033	
5% level	-1.947520	
10% level	-1.612650	

\*Mackinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation  
 Dependent Variable: D(DTREUSSIT)  
 Method: Least Squares  
 Date: 01/03/23 Time: 14:20  
 Sample (adjusted): 1972 2021  
 Included observations: 50 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
DTREUSSIT(-1)	-2.552770	0.383198	-6.661759	0.0000
D(DTREUSSIT(-1))	1.236162	0.306610	4.031702	0.0002
D(DTREUSSIT(-2))	0.648548	0.218431	2.969115	0.0047
D(DTREUSSIT(-3))	0.343683	0.137017	2.508322	0.0157
R-squared	0.684501	Mean dependent var	-0.100000	
Adjusted R-squared	0.663925	S.D. dependent var	18.55081	
S.E. of regression	10.75427	Akaike info criterion	7.665102	
Sum squared resid	5320.102	Schwarz criterion	7.818064	
Log likelihood	-187.6275	Hannan-Quinn criter.	7.723351	
Durbin-Watson stat	1.954170			

Table 5: Application of the ADF Test to the first differentiation series

Source: Author using eviews 12 software

We now have a prob= $\alpha < 5\%$ ; the ADF value = -6.661759 which is less than the Mackinnon critical value equal to -1.947520. We accept  $H_1$  and we reject  $H_0$ . Hence the TREUSSIT series is stationary after the first difference.

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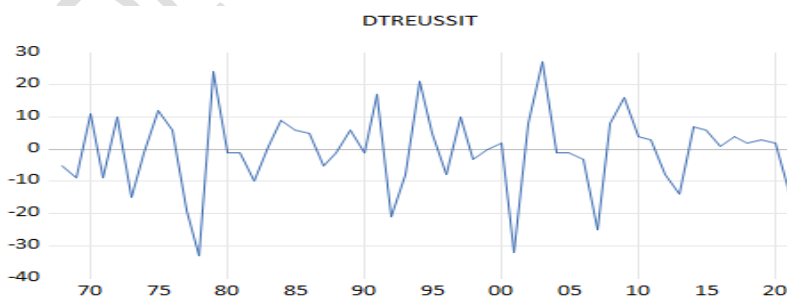


Figure 3: Graph of the stationary TREUSSIT series

Source: Author using eviews 12 software

We notice that there are fluctuations around the mean, which shows that our series is stationary. From where we can move on to the step of determining p and q.

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## 4.2. Modelization

### 4.2.1 Determination of orders (p, q)

We have a stationary series, we now need to determine an ARMA model (p, q); To do this, we will use the correlogram from the stationary TREUSSIT series. The simple correlogram allows us to identify an MA(q) model, and the partial correlogram allows us to identify an AR(p) model.

Please note that it is advisable not to exceed the first three delays so as not to lose a lot of information.

Here is the correlogram of the stationary TREUSSIT series after the first difference;

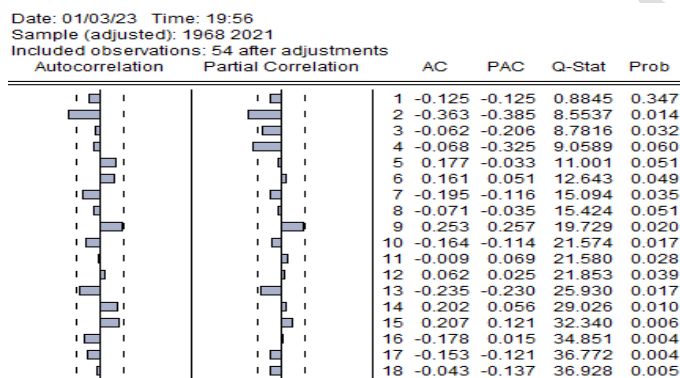


Table 6: Correlogram of the stationary TREUSSIT series

Source: Author using eviews 12 software

We note that certain probabilities of the correlogram of the differentiated series are less than 5%, it is therefore not a question of a random approach, that is to say that there exists a representation of the series in the family of the ARMA model.

By differentiating our series, we removed the trend it presented in order to stationarize it. The ARIMA model is therefore a combination of this differentiation process and the classic ARMA process. Furthermore, if the time series presents both a trend and seasonality, it is possible to use the SARIMA model, which is nothing more than an ARIMA model taking into account a seasonal component. We offer the following models with constant and order 1 differentiation:

- ARIMA (2, 1,2)
- ARIMA (0, 1,2)
- ARIMA (2, 1,0)

The models having been proposed, we can move on to the stage of estimating its parameters.

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### 4.3. MODEL ESTIMATION

The model is estimated with the stationary series; we will accept the model with significant coefficients; if the model parameters are significant then we will be able to test these residuals which must be white noise.

We will have to compare the values found to the ratio of Student's *t* statistic (for us 1.96). Using the Eviews 12 software. We have the following estimates:

#### 4.3.1 Estimation of the ARIMA model (2, 1, 0)

Dependent Variable: DTREUSSIT  
Method: ARMA Conditional Least Squares (Gauss-Newton / Marquardt steps)  
Date: 01/03/23 Time: 14:30  
Sample (adjusted): 1970 2021  
Included observations: 52 after adjustments  
Convergence achieved after 9 iterations  
Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.028017	1.175798	0.023828	0.9811
AR(2)	-0.372776	0.132147	-2.820918	0.0069

R-squared	0.137300	Mean dependent var	0.038462
Adjusted R-squared	0.120046	S.D. dependent var	12.40803
S.E. of regression	11.63946	Akaike info criterion	7.784382
Sum squared resid	6773.853	Schwarz criterion	7.859430
Log likelihood	-200.3939	Hannan-Quinn criter.	7.813153
F-statistic	7.957580	Durbin-Watson stat	2.445264
Prob(F-statistic)	0.006850		

Inverted AR Roots	-0.00+0.61i	-0.00-0.61i
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Table 7: Estimation of the ARIMA model (2, 1, 0)

Source: Author using eviews 12 software

We have the probability linked to AR(2) equal to 0.0069 which is less than 0.05 (5%); the t-statistic value is greater than 1.96; this shows that the coefficient of AR(2) is significantly different from zero. Hence the model is good and is retained for the test of the residuals.

#### 4.3.2 Estimation of the ARIMA model (0, 1, 2)

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Dependent Variable: DTREUSSIT  
 Method: ARMA Conditional Least Squares (Gauss-Newton / Marquardt steps)  
 Date: 01/03/23 Time: 14:32  
 Sample (adjusted): 1968 2021  
 Included observations: 54 after adjustments  
 Failure to improve likelihood (non-zero gradients) after 15 iterations  
 Coefficient covariance computed using outer product of gradients  
 MA Backcast: 1966 1967

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.026015	0.742054	0.035058	0.9722
MA(2)	-0.534244	0.113742	-4.696989	0.0000
R-squared	0.189632	Mean dependent var		-0.222222
Adjusted R-squared	0.174048	S.D. dependent var		12.25156
S.E. of regression	11.13445	Akaike info criterion		7.694298
Sum squared resid	6446.749	Schwarz criterion		7.767964
Log likelihood	-205.7461	Hannan-Quinn criter.		7.722708
F-statistic	12.16836	Durbin-Watson stat		2.560588
Prob(F-statistic)	0.000998			
Inverted MA Roots	.73	-.73		

Table 8 : Estimation of the ARIMA model (0, 1,2)

Source: Author using eviews 12 software

We have the probability linked to the MA(2) model, prob=0, less than 0.05 and the coefficient of the model significantly differs from zero because the t-statistic value is well above 1.96; we can therefore retain the model for testing the residuals.

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#### 4.3.4 Estimation of the ARIMA model (2, 1,2)

Dependent Variable: DTREUSSIT  
 Method: ARMA Conditional Least Squares (Gauss-Newton / Marquardt steps)  
 Date: 01/03/23 Time: 14:27  
 Sample (adjusted): 1970 2021  
 Included observations: 52 after adjustments  
 Failure to improve likelihood (non-zero gradients) after 26 iterations  
 Coefficient covariance computed using outer product of gradients  
 MA Backcast: 1968 1969

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.160583	0.695442	0.230908	0.8183
AR(2)	0.132592	0.236241	0.561259	0.5772
MA(2)	-0.640104	0.186311	-3.435680	0.0012
R-squared	0.227966	Mean dependent var		0.038462
Adjusted R-squared	0.196454	S.D. dependent var		12.40803
S.E. of regression	11.12265	Akaike info criterion		7.711805
Sum squared resid	6061.951	Schwarz criterion		7.824377
Log likelihood	-197.5069	Hannan-Quinn criter.		7.754962
F-statistic	7.234355	Durbin-Watson stat		2.488376
Prob(F-statistic)	0.001766			
Inverted AR Roots	.36	-.36		
Inverted MA Roots	.80	-.80		

Table 9 : Estimation of the ARIMA model (2, 1,2)

Source: Author using eviews 12 software

On the one hand, we have the probability of AR(2) greater than 5%; that of MA(2) less than 5% and on the other side, the coefficient of the AR(2) process significantly equal to zero because its t-statistic value is less than 1.96 and that of MA(2) significantly

differs from zero because its t-statistic value is much greater than 1.96. Hence the model overall is not good and should be rejected

#### 4.4 Validation

##### 5.1 Residue testing

The estimation phase allowed us to retain two models that we must test their residuals to find the model that would contain white noise residuals. Hence the residue test.

##### 4.4.1 Test for autocorrelation of residuals

- a) Testing the residuals of the ARIMA model (2, 1.0)

##### Residual correlogram

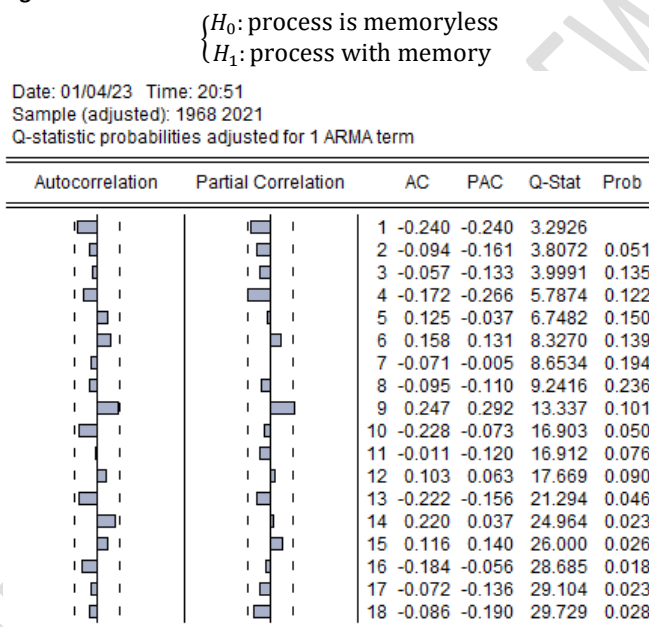


Table 10 : Residual correlogram of the ARIMA model (2, 1.0)

Source: Author using eviews 12 software

We note that the probabilities are greater than 5% up to order  $k > 13$ ; from which we can say that we have a process whose residues do not follow the course of chance. Thus, we resort to the statistical test of  $Q_{BP}$ .

Consider the Box-perce statistic at delay  $k=18$ ,  $Q_{BP} = 29.729$  to compare with the value of  $\chi^2$  at delay  $kq-1$  is equivalent to  $k=15$ ;  $\chi^2=24.99$  which is lower than the statistical value  $Q_{BP}$ .

So this is not a memoryless process; Hence the model is to be rejected.

- b) Test of residuals on the ARIMA model (0, 1,2)

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## Residual correlogram

$$\begin{cases} H_0: \text{process is memoryless} \\ H_1: \text{process with memory} \end{cases}$$

Date: 01/04/23 Time: 14:54  
 Sample (adjusted): 1968 2021  
 Q-statistic probabilities adjusted for 1 ARMA term

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob	
		1	-0.278	-0.278	4.4142	
		2	0.009	-0.075	4.4184	0.036
		3	-0.131	-0.162	5.4297	0.066
		4	-0.016	-0.112	5.4448	0.142
		5	0.098	0.051	6.0323	0.197
		6	0.122	0.161	6.9715	0.223
		7	-0.070	0.015	7.2900	0.295
		8	-0.092	-0.083	7.8480	0.346
		9	0.227	0.252	11.314	0.185
		10	-0.185	-0.075	13.668	0.135
		11	0.000	-0.142	13.668	0.189
		12	0.053	0.068	13.874	0.240
		13	-0.189	-0.194	16.500	0.169
		14	0.163	-0.003	18.515	0.139
		15	0.120	0.185	19.629	0.142
		16	-0.174	-0.106	22.031	0.107
		17	-0.064	-0.119	22.366	0.132
		18	-0.116	-0.194	23.502	0.134

Table 11: ARIMA residual correlogram (0, 1, 2)

Source: Author using Eviews 12 software

We note that the simple and partial autocorrelations are all significantly equal to zero because its values do not leave the Distaff band at the 5% threshold, with the probabilities all greater than 0.05 except at lags k=2.

To do this, we move on to the statistics  $Q_{BP}$  of the series which we will compare with the value of  $\chi^2$  at delay 15 and we notice that  $\chi^2 > Q_{BP}$ , at the threshold of 5%. That is to say  $24.99 > 23.50$ . Hence, we assimilate the process to that of chance and accept  $H_0$  because it is a process without memory, which means that it will not intervene in the forecasts.

### 4.4.2 Test of Heteroskedasticity

#### a) Arch test

$$\begin{cases} H_0: \text{the residuals are not autocorrelated (presence of homoscedasticity)} \\ H_1: \text{the residuals are autocorrelated (presence of heteroscedasticity)} \end{cases}$$

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Heteroskedasticity Test: ARCH

F-statistic	0.330381	Prob. F(2,49)	0.7202
Obs*R-squared	0.691887	Prob. Chi-Square(2)	0.7076

Test Equation:

Dependent Variable: RESID^2

Method: Least Squares

Date: 01/03/23 Time: 14:42

Sample (adjusted): 1970 2021

Included observations: 52 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	110.5092	35.59694	3.104458	0.0032
RESID^2(-1)	0.115890	0.142793	0.811596	0.4209
RESID^2(-2)	-0.007312	0.142840	-0.051187	0.9594
R-squared	0.013306	Mean dependent var		123.8104
Adjusted R-squared	-0.026968	S.D. dependent var		191.8767
S.E. of regression	194.4467	Akaike info criterion		13.43415
Sum squared resid	1852666.	Schwarz criterion		13.54673
Log likelihood	-346.2880	Hannan-Quinn criter.		13.47731
F-statistic	0.330381	Durbin-Watson stat		2.003278
Prob(F-statistic)	0.720238			

Table 12: Heteroscedasticity test (ARCH) on the ARIMA model (0, 1,2)

Source: Author using Eviews 12 software

We have the test probability of ARCH = 0.7 which is greater than 5%, The statistical value of Durbin-Watson (DW); DW = 2.003278  $\approx$  2 which shows the absence of autocorrelation of the residuals. Hence, we accept  $H_0$  the absence of Heteroskedasticity, which is also Demonstrated by the correlogram below:

b) Correlogram of the squares of the residuals

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Date: 01/04/23 Time: 16:01  
Sample (adjusted): 1968 2021  
Included observations: 54 after adjustments

	Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1	0.117	0.117	0.7873	0.375		
2	0.007	-0.007	0.7898	0.674		
3	-0.160	-0.162	2.3072	0.511		
4	0.011	0.050	2.3141	0.678		
5	-0.011	-0.016	2.3214	0.803		
6	0.168	0.150	4.0958	0.664		
7	-0.076	-0.112	4.4638	0.725		
8	-0.206	-0.206	7.2665	0.508		
9	0.063	0.189	7.5336	0.582		
10	0.034	-0.031	7.6146	0.666		
11	-0.138	-0.231	8.9444	0.627		
12	-0.031	0.051	9.0115	0.702		
13	0.015	0.059	9.0270	0.771		
14	0.032	0.035	9.1029	0.824		
15	0.014	-0.093	9.1188	0.871		
16	-0.012	-0.052	9.1302	0.908		
17	-0.176	-0.022	11.667	0.820		
18	-0.087	-0.113	12.297	0.832		

table 13 : Correlogram of the squares of the residuals of the ARIMA model (0.1, 2)

Source: Author using Eviews 12 software

The correlogram presents autocorrelation coefficients all significantly equal to zero at the 5% threshold because none of the simple or partial autocorrelation coefficients leave the distaff band.

**With probabilities greater than 5%, we can say that there is the absence of autocorrelation of the residuals which justifies the presence of homoscedasticity of the errors.**

#### 4.4.3 Normality test

$H_0$ : the residuals are not Gaussian white noise  
 $H_1$ : the residuals are Gaussian white noise

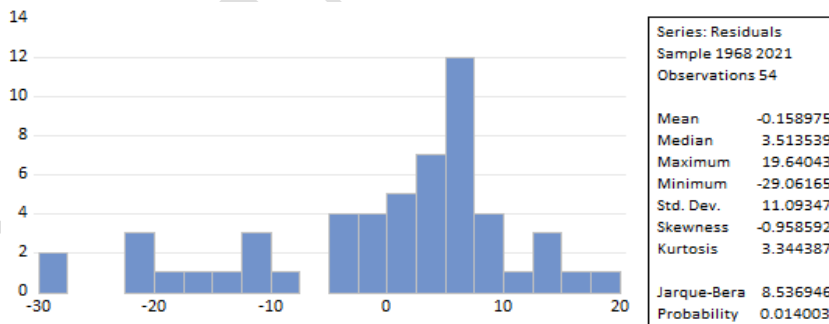


Figure 4: The histogram of the normality test of the residuals (Jacques Berat test)

Source: author using eviews 12 software

We see from the histogram of the residuals that:

The skewness value being -0.958592 is less than 1.96 therefore the residuals are distributed symmetrically.

That of kurtosis is  $3.344387 > 1.96$  explains that the distribution of residuals is not flattened and finally that of the Jacques-Berat statistic is 8.536946 which is  $> \alpha 0,05$  at the threshold value ;

Which explains why the residues follow a normal law; from which we accept  $H_1$  and reject  $H_0$ , the residuals are therefore Gaussian white noise.

It should be noted that this process is memoryless and there is no autocorrelation of the residuals; we can say that the residuals are Gaussian white noise because the assumption of normality of the residuals is respected.

We retained the ARIMA model (0, 1,2) because it is the only one in the ARIMA process to have satisfied the conditions required to forecast with the Box-Jenkins method.

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### 5. Forecast

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To make forecasts, it is necessary to modify the size of the series, up to the forecast period envisaged. We know that the residuals are white noise; that is to say they cannot intervene in the forecast, we should therefore use the forecast formula of an ARIMA model given by:

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$$DTREUSSIT. F_t = c + \sum_{i=1}^{p+d} \theta_i X_{t-i} + \varepsilon_t + \sum_{k=j}^q \varphi_j \varepsilon_{t-j}$$

And

$$\hat{X}_{n+h} = c + \sum_{i=1}^{p+d} \theta_i \hat{X}_{n+h-i} + \sum_{j=1}^q \varphi_j \hat{\varepsilon}_{n+h-j}$$

$$\text{If } k \geq h \text{ and } \hat{\varepsilon}_{n+h-k} = \begin{cases} \varepsilon_{n+h-k} & \text{si } k \geq h \\ 0 & \text{si } k < h \end{cases}$$

If we replace the data we have in the formula, we will have;

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For h=1

$$X_t = c + \sum_{k=1}^{0+1} \theta_k X_{t-k} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k}$$

$X_t = c + \sum_{k=1}^{0+1} \theta_k X_{t-k} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k}$  Which brings us to;

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$$X_t = c + \theta_1 X_{t-1} + \varepsilon_t + \sum_{k=1}^2 \varphi_k \varepsilon_{t-k} \text{ And}$$

$$\hat{X}_{55+1} = \sum_{k=1}^{0+1} \theta_k \hat{X}_{55+1-k} + c + \sum_{k=1}^2 \varphi_k \hat{\varepsilon}_{55+1-k}$$

$$\hat{X}_{56} = \theta_1 \hat{X}_{55} + c + \sum_{k=1}^2 \varphi_k \hat{\varepsilon}_{55-k}$$

With  $h \geq q$ ; we obtain a relation of the form;  $\hat{X}_{n+h} = c + \sum_{k=1}^{p+d} \theta_k \hat{X}_{n+h-k}$

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For h=2,

$$\hat{X}_{57} = c + \sum_{k=1}^{0+1} \theta_k \hat{X}_{55+2-k}$$

$$\hat{X}_{57} = c + \theta_1 \hat{X}_{55+2-1}$$

$$\hat{X}_{57} = c + \theta_1 \hat{X}_{56}$$

Of this ; we can deduce;

For h=3,

$$\hat{X}_{58} = c + \theta_1 \hat{X}_{57}$$

For h=4,

$$\hat{X}_{59} = c + \theta_1 \hat{X}_{58}$$

We know that the results above are estimated from Dtreussit because  $DTREUSSIT_t = TREUSSIT_t - TREUSSIT_{t-1}$

$$SO ; TREUSSITF_t = DTREUSSITF_t + TREUSSITF_{t-1}$$

$$\text{If } h=1, \text{ then } TREUSSITF_{t-1} = TREUSSIT_{t-1}$$

We used the Eviews 12 software to have the following forecasts for the years 2022, 2023, 2024, 2025:

Table 14: Forecasts up to h=4

YEAR	TREUSSITF
2022	56.17347813585213
2023	61.6376834440594
2024	61.60563624509442
2025	61.57358904612945

These forecasts are also given by the following diagram

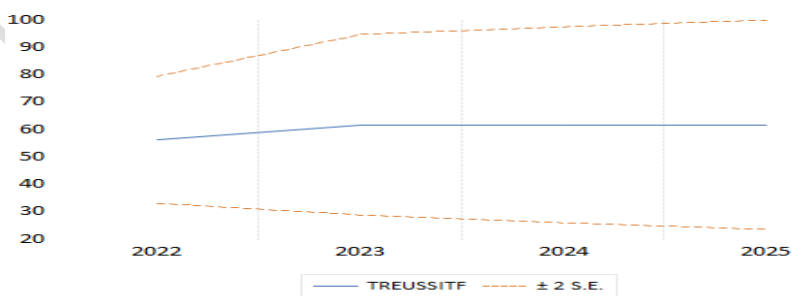


Figure 5: Graph of predicted values  
Source: author using eviews 12 software

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The predicted values of the TREUSSIT series are obtained after determining the errors observed by the moving average of the differentiated series, it is therefore a short memory model. As a result, we can only make short-term forecasts.

The red lines represent the two limits of the confidence interval which appears to be constant from 2023 until 2025.

So, here is the graphic representation of the treussit series until the year 2025

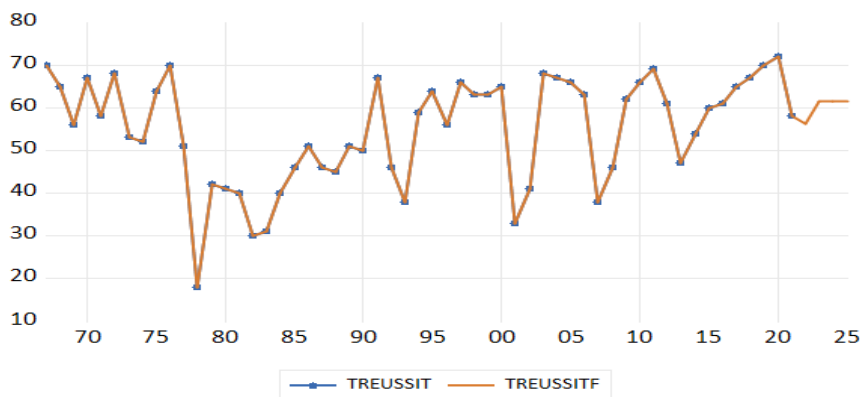


Figure 6: Graphical representation of the Treussit series predicted up to the value  $h=4$

Source: author using eviews 12 software

## 6. Discussion :

Our study focuses on the analysis of the variable “success rate for State Examinations in the DRC” which means studying the evolution of these rates and predicting in the short term as the model requested. It should be noted that these rates are influenced by several external variables. Let us say in passing that the Congolese education system is shaken by repeated strikes especially at the start of each school year, the mobility of students who fail before reaching the last year or even the last year; constant wars in the east of the country; the change of school programs without training of trainers and without documentation; unsatisfied motivation of trainers; unsuitable frameworks for training; the politicization of the education system; the lack of working tools on the part of both learners and trainers; The epidemics ; corruption and anti-values of all kinds; ...

All these factors contribute in one way to another to vary these rates.

In developing these forecasts, we have not taken into account the effects of these factors on the pass rates for the state exam in the DRC. Our only factor is time. Which justifies the name time series.

The most parsimonious model is given after first differentiation; for our TREUSSIT series; we retained the ARIMA model (0, 1,2), which allowed the prediction.

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## CONCLUSION

The main goal was to predict. The model imposed on us short-term forecasts of these success rates.

To achieve this, we chose the Box-Jenkins approach, a parsimony approach. We used a computer tool which is the Eviews 12 software, which made it possible to obtain the desired results.

After the presentation of success rate data, we moved on to the graphical representation which constitutes the first step of the Box-Jenkins methodology; this graph did not completely confirm to us that the series was non-stationary; the study of the correlogram of partial and simple autocorrelation functions, the augmented Dickey Fuller test (ADF) showed the non-stationarity of this so-called TREUSSIT series and the way of stationarizing it.

Results obtained after testing the ADF for its models were presented as follows:

- Model 3, test on the constant and the trend gave the absence of the trend and the presence of the constant,
- Model 2; test on the constant indicates the presence of the constant,
- Model 1; test without constant or trend affirmed that there is the presence of a unit root.

And these results led us to a non-stationary model of the DS type with drift; the series becomes stationary after applying the first difference. Hence the TREUSSIT series becomes DTREUSSIT.

Three models are proposed ARIMA(2,1,0); ARIMA(0,1,2); ARIMA(2,1,2), among these three, only the ARIMA(0,1,2) model turned out to be the only one to have white noise and Gaussian residuals and was retained for calculating the forecasts of the series DETROIT on the horizons h.

The short-term forecasts calculated using the eviews12 software are: 56.17% for 2022, 61.64% in 2023, 61.61% in 2024 and 61.57% in 2025.

To show the veracity of our results, we invite readers to compare them with data actually recorded in the field. Already, for the year 2022, the success rate is 60%, slightly lower than that proposed by our model.

The forecasts obtained are very important for a view on the future of success rates in the State Examination which is an instrument for evaluating National Education at the end of the secondary cycle. However, we say that several factors can influence these forecasts. We cited natural disasters, pandemics, free education, anti-values,...

We believe that the Congolese State will take its results into account to improve educational results overall.

We propose the following themes for the future:

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- ✓ Time series analysis of the results of each study cycle in the DRC
- ✓ Comparative study of the series of State Examination results by Province in the DRC;
- ✓ Impact of anti-values on Exeat success rates in the DRC

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