

## EMPIRICAL INVESTIGATION ON THE EFFECT OF THE NUMBER OF RESAMPLINGS ON THE DISTRIBUTION OF BOOTSTRAP STANDARD ERROR USING RESPONSE TIME DATA

### ABSTRACTS.

**Aims:** To investigate how the number of bootstrapping  $B$  affects the values returned by the bootstrap standard error of the arithmetic mean and the  $\alpha$ -trimmed mean of response data using bootstrap confidence intervals (CI) at 95% level. Carried out to fill up observed gap for study on the standard error (the tool generally employed in assessing the long run accuracy of a given statistical estimator of  $\theta$ ) of response data.

**Study design:** A non-parametric, empirical bootstrap simulation study.

**Place and Duration of the study:** Two departments: Computer Science and Statistics, Federal Polytechnic Oko, Anambra state Nigeria 2020/2021 session.

**Methodology:** Response time data were generated with student customers of mobile telephone network (mtn) Nigeria and stored in SPSS. A sample  $n = 51$  responses were selected using "Select Cases" command to increase precision and minimize bias. Bootstrap simulation study was carried out using R programming language. Four approaches for estimating bootstrap confidence intervals were used. The interval coverage and the interval lengths were determined and compared for  $B = 20, 50, 100, 500, 1000, 5000,$  and  $10000$ .

**Results:** The 95% CI for  $s_{10} = 0.2266338$  (the estimated sample standard error of 10% trimmed mean) returned the best interval for our skewed data set; when  $B = 20$ ; the CI for  $s_{10}$  returned  $(0.2051, 0.2343)$  for the normal approach,  $(0.2082, 0.2371)$  for the basic,  $(0.2071, 0.2360)$  for the percentile, and  $(0.2071, 0.2360)$  for the BCa method. As  $B$  increased to 5000, it returned  $(0.2259, 0.2277)$  for the normal approach,  $(0.2259, 0.2277)$  for the basic,  $(0.2260, 0.2277)$  for the percentile, and  $(0.2261, 0.2279)$  for the BCa showing a shorter interval yet covering the estimate.

**Conclusion:** Thus for our response data study, increasing  $B$  in estimating standard error increases the chances of more precise and shorter confidence intervals rather than the chances for coverage.

**Keywords:** *Bootstrap, standard error, trimmed mean, number of bootstrapping, response data*

### 1 INTRODUCTION:

A bootstrap method is a resampling technique often employed in the construction of confidence intervals to estimate the variation of point estimates. The approach is largely based on the law of large numbers (Orloff, and Bloom, 2017) [1]. An article published by John Wiley and sons Ltd Dixon (2006) [2] noted that at the heart of the bootstrap method is the concept that the distribution of the statistic of interest

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can be approximated by estimates from repeated samples. The bootstrap technique makes use of the information available in the data without making any ad hoc assumptions about the distribution of the population of the data. The bootstrap is a general methodology for answering questions on the accuracy of an estimator of  $\theta$  and the standard error is often used to measure long run accuracy of such estimator. A concern however is often on the number of resampling that will yield the optimum estimate of interest especially for asymmetric data set. Much work has been done on the effect of the number of resampling on the distribution of the bootstrap mean and other measures of the center but not so much on the distribution of the bootstrap standard error.

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The distributions of response time (RT) data violate the normality assumption underlying classical statistical analyses because they are generally positively skewed and unimodal in shape (Heathcote, Popiel, and Mewhort, 1991)[3]. Customers of service providers in Nigeria of which the researcher is one, often experience delays from the service providers in responding to their online complaints via their online customer care. This can be frustrating most times. Service providers' response to online customers' complaints is a response time data set and as such should be expected to be skewed.

According to Chernick and LaBudde (2011)[4], the arithmetic sample mean is the theoretical mean of the bootstrap distribution. Lampert, Stahel, and Abbt (2001)[5] opined that response time data distributions are mostly skewed and the arithmetic mean generally delivers a poor measure of central tendency for skewed distributions. The  $\alpha$ -trimmed mean provides more accurate information about the central tendency than the mean for an asymmetric distribution. Since the sample mean is the theoretical mean of the bootstrap distribution and the  $\alpha$ -trimmed mean is the preferred mean for skewed data, we decided to study on the bootstrap standard error of the mean as well as that of the  $\alpha$ -trimmed mean.

Our interest was to investigate how the number of resampling affects the values returned by the bootstrap standard error of the arithmetic mean and that of the  $\alpha$ -trimmed mean using bootstrap confidence interval at 95% level. Researchers sometimes spend much resources trying to run large number of resampling in setting up bootstrap confidence intervals hoping possibly to increase the chances of the interval including the estimate of interest.

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Writing on the needed number of bootstrap samples, Davidson and Mackinnon (2001)[6], Orloff and Bloom (2017)[1] believe that the number of bootstrap resampling can influence the accuracy of the bootstrap estimate. Specifically in the opinion of Davidson and Mackinnon (2001)[6], where the cost of computation is affordable,  $B$  should be extremely large.

Four different approaches are considered to estimating bootstrap confidence interval namely: the first order normal approximation, the basic bootstrap interval, the bootstrap percentile interval, and the adjusted bootstrap percentile that is the bias-corrected accelerated (BCa) interval. These were used to set up the confidence intervals for the selected measures. How these different approaches behave in the estimation of the confidence interval of the standard error for the mean and trimmed mean for skewed response time data and as the number of bootstraps increased was considered.

Efron and Tibshirani (1986) [7] opined that for estimating the standard errors and coefficient of Variations (CV), B as small as 25 gives reasonable results. They iterated however that the situation is different for the construction of confidence intervals. Efron and Tibshirani [7], quoting the calculations of Efron (1984) [8], they stated that B = 1000 is a rough minimum for the number of Monte Carlo bootstraps necessary to compute BC or BCa. They maintained however that somewhat smaller values like B = 250 can give a useful percentile interval. Davison and Hinkley (1997) [9] recommended the number of bootstraps B = 999 for confidence interval levels of 95% and 99% where such is practically feasible.

Using computer packages today, it will take a few seconds to reach such number. The question however is how relevant is it to keep B increasing?

From the foregoing, therefore, it is expected that the more the number of resampling B, the more accurate the estimate and the shorter the confidence interval. It is therefore expected that the interval will be getting shorter as B gets larger and with an increased probability of coverage. Using the service provider data therefore and employing the confidence interval tool, we studied how B affects the accuracy and the precision of the estimates of bootstrap standard error for the selected bootstrap methods for calculating CI at 95% level.

The bootstrap standard error of the arithmetic means as well as that of the  $\alpha$ -trimmed mean and their 95% confidence intervals were estimated using the four approaches and compared. The interval lengths were equally calculated and compared. The study covered for B (the number of bootstrap resampling) = 20, 50, 100, 500, 1000, 5000, 10000.

## 2. MATERIALS AND METHODS:

### 2.1 Data Collection Procedure

The data for this research was obtained with from Statistics and Computer Science students of Federal Polytechnic Oko Atani campus, Anambra state Nigeria. The selected students were all customers of Mobile Telephone Network (MTN) Nigeria; one of the major service providers in the south-east of Nigeria.

Volunteer participants were utilized to obtain the data. In the course of obtaining the data, the participants were asked to place calls to their network provider MTN using the network's call center number at separate times. The intention of a caller/customer in using the online customer care is always to speak with the online customer care representative within the shortest possible time. Often times, however, when calls are placed on them, some messages are being played in the form of adverts and jingles before transferring the caller to a customer care representative. Some network providers will display their adverts and jingles for a long time before finally allowing the caller/customer to speak with an agent. This could be very frustrating as well as time and money-consuming. I personally do not like that and it often discourages customers from utilizing such services.

The participants were told to call their customer care line and follow the required processes until they are transferred to a customer care agent to attend to them. They were to end the call immediately after the

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customer care representative responds to the call and record the total time spent in minutes. These were done severally for different periods and the time/duration of the calls recorded. At the end, the data were retrieved, organized, and entered in the Statistical Product and Service Solutions (Hejase and Hejase, 2013 [10] IBM SPSS and the desired sample was randomly selected using the "Select Cases" command from the data menu with the purpose to increase precision and avoid bias.

## 2.2 Theory and Methods:

### 2.2.1 The Bootstrap Principle and Procedure:

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The bootstrap principle believes that the empirical distribution obtained by resampling from a sample is approximately equal to the theoretical or true distribution (Orloff, and Bloom; 2017)[1].

Let  $\theta$  = any statistics computed from the original sample data;

$\theta^*$  = a statistic by the same formula as  $\theta$  but computed instead using resampled data from the original sample.

To calculate the desired bootstrap estimate  $\theta^*$ , we apply the following procedure:

- (1) Select a sample of size  $n; x_1, x_2, x_3, \dots, x_n$  from a distribution  $F$  to form an empirical distribution  $F_n$  which puts equal mass  $\frac{1}{n}$  on the selected  $n$  data points.
- (2) Calculate the desired estimate  $\theta$ .
- (3) Select a sample of size  $n = x_1^*, x_2^*, x_3^*, \dots, x_n^*$  with replacement from the original sample in (1) above
- (4) Calculate  $\theta_i^*$  the value of  $\theta$  from the  $i$ th resampled data using the same formula for calculating  $\theta$ .
- (5) Repeat (3) and (4)  $B$  times so that we have  $\theta_1^*, \theta_2^*, \theta_3^*, \dots, \theta_B^*$
- (6) The bootstrap estimate of  $\theta$ ,  $\theta_b$ , is obtained by calculating  $\theta$  still using the same formula but based on  $\theta_1^*, \theta_2^*, \theta_3^*, \dots, \theta_B^*$ . (Wilcox, 2005)[11]

### 2.2.2 Bootstrap Estimate of the Standard Error of the Mean:

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The long-run accuracy of a given statistical estimator of  $\theta$  is usually assessed with its standard error. (Staudte and Sheather, 1990)[12]

The standard error of the arithmetic mean given  $n$  samples is given as

$$s_x = \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}} \quad (\text{Madsen and Moeschberger; 1986}) [13]$$

Following the procedure for bootstrap estimates presented above, the bootstrap estimate for the standard

error of the standard error of arithmetic mean is given as:  $s_b = \sqrt{\frac{\sum_{i=1}^B (s_i^* - \bar{s}^*)^2}{B-1}}$ :

(1)

where  $s_b$  = the bootstrap standard error of the standard error.

$s_i^*$  = the standard error of the  $i$ th bootstrap sample,

$$\bar{s}^* = \frac{\sum_{i=1}^B s_i^*}{B} = \text{the arithmetic mean of the } B \text{ bootstrap sample}$$

standard errors and  $B$  = the number of bootstrap samples.

### 2.2.3 Bootstrap Estimate of the Standard Error of the $\alpha$ -Trimmed Mean.

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The standard error of the  $\alpha$ -trimmed mean is given as:  $s_{\alpha} = \sqrt{\frac{s_w^2}{(1-\alpha)^2 n}} = \frac{s_w}{(1-\alpha)\sqrt{n}}$  (2)

Where  $s_w^2 = \frac{1}{n-\alpha} \sum_{i=1}^n (w_i - \bar{w})^2$  = sample winsorized variance (Staudte and Sheather, 1990, [12]

Wilcox; 2005 [11])

The bootstrap version  $s_{b\alpha} = \sqrt{\frac{s_{bw}^2}{(1-\alpha)^2 B}} = \frac{s_{bw}}{(1-\alpha)\sqrt{B}}$ ; where  $s_{bw}^2 = \frac{1}{B-\alpha} \sum_{i=1}^B (w_i - \bar{w})^2$  (3)

The limit cases of  $\alpha = 0$  and  $\alpha = 0.5$  correspond to the sample mean and median respectively (Staudte and Sheather, 1990 [12], Maronna et al, 2006 [14]). The choice of  $\alpha$  is of practical importance. It is generally chosen so that  $\bar{x}_{\alpha}$  is apt to have a relatively small standard error  $\sqrt{\text{var}(\bar{x}_{\alpha})}$  among commonly occurring situations.

Wilcox (2005) [11] noted that if  $\alpha$  is too small, the standard error of the trimmed mean can be drastically inflated by outliers or by sampling from a heavy-tailed distribution. When one is sampling from a normal distribution if the percentage of trimming  $\alpha$  is quite large, the standard error of the trimmed mean can be relatively large when compared to that of the arithmetic sample mean.

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### 2.3 Method of Comparison.

The method of comparison employed for the study is confidence interval (CI). CI is used to measure the precision of an estimate, to know the plausible range of values given the parameter of interest. It is the probability that the computed interval will include the population parameter (Efron & Tibshirani, 1993) [15]. There are several available approaches for estimating bootstrap confidence intervals but as earlier stated, for this study four methods were considered. These are the first-order normal approximation, the basic bootstrap interval, the bootstrap percentile interval, and the adjusted bootstrap percentile that which is the bias-corrected accelerated (BCa) interval (Canty, 2021) [16].

The first-order normal approximation is calculated using the normal approximation. The basic bootstrap interval uses the estimated standard error and is calculated using the basic bootstrap method, the bootstrap percentile interval and the BCa intervals use the bootstrap percentile method but the BCa interval in addition is also adjusted to account for bias and skewness (Larget, 2014) [17].

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A 95% confidence interval was set up using the four approaches named above.

We compared the confidence interval coverage of the selected procedures for  $B$ , (the number of bootstrap resampling) = 20, 50, 100, 500, 1000, 5000, and 10000.

The length of each confidence interval was determined as:

$$\text{length} = \hat{\theta}_{bup} - \hat{\theta}_{blo} \text{ where } \hat{\theta}_{bup} = \text{upper value of the bootstrap interval} \\ \hat{\theta}_{blo} = \text{the lower value of the bootstrap interval}$$

John Wiley and Sons Ltd in their article published in *Encyclopedia of Environmetrics* Dixon (2006)[2] opined that a percentile interval may not have the correct coverage for a skewed sampling distribution. Efron (1984)[8] showed that B=1000 is a rough minimum for the number of Monte Carlo bootstraps needed to compute BCa intervals. Efron (1987)[18] stated that the BCa approach to calculating the bootstrap confidence interval properties that puts it at advantage over the other methods. The coverage errors for the confidence interval for the mean and the median is usually small using the BCa method whether the population distribution is normal or non-normal even for a sample size of about 20 or more. (Efron & Tibshirani, 1993[15]; Lei & Smith, 2003[19]).

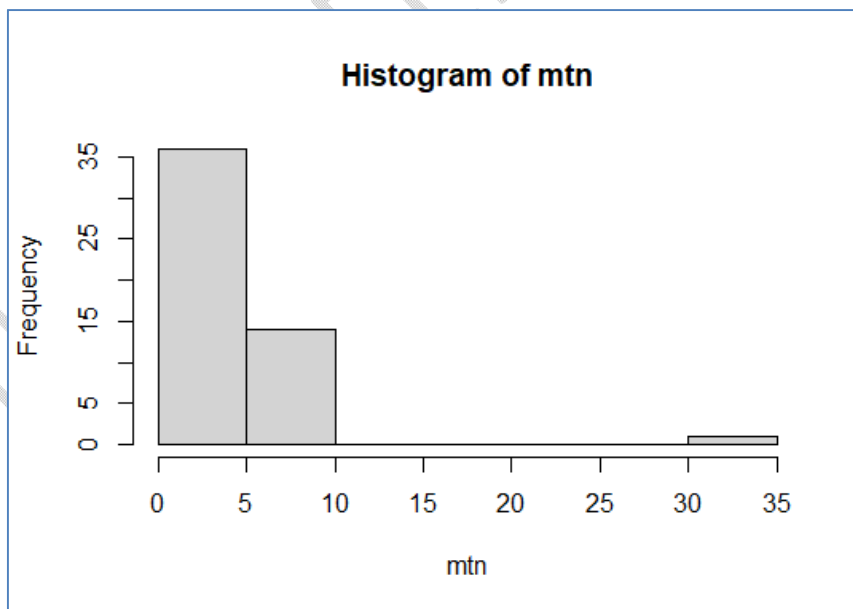
Setting up a (1- $\alpha$ )% confidence interval implies that on the average one would expect (1- $\alpha$ )% of the intervals to include the true value of the estimate while  $\alpha$ % would not include it (Glasserman, 2001[20], Helwig, 2017[21]).

The R programming language was employed both for the simulation and analysis.

### 3. RESULTS AND DISCUSSIONS.

The data for this research work is presented on a histogram below. As should be expected from our literature review, the data is positively skewed being response data (Figure 1).

**Fig 1. Histogram of the Response Data Used for the Study**



**Fig 1. Histogram of the Response Data Used for the Study**

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From the analysis, the estimated sample standard error  $s_x$  of the arithmetic mean for the  $n=51$  set of data studied was 4.123277, that of 10% trimmed mean  $s_{10}$  was 0.2266338, and  $s_{20} = 0.2663485$  for the sample standard error of 20% trimmed mean. Thus, the  $\alpha$ -trimmed mean returned smaller standard error than the arithmetic mean with the 10% trimming returning slightly smaller value than 20% trimming this could be possibly explained by the non-symmetric nature of our data set.

Table 1 is a summary of the results of the analysis while Table 2 is a display of the width of the bootstrap confidence intervals.

**Table 1: 95% Confidence Intervals for the Bootstrap standard Errors ( $s_x$ ,  $s_{10}$  and  $s_{20}$ )**

B	Standard Error of the Mean ( $s_x$ )				Standard Error of the 10% Trimmed Mean ( $s_{10}$ )				Standard Error of the 20% Trimmed Mean ( $s_{20}$ )			
	FONA	BBI	BPI	BCa	FONA	BBI	BPI	BCa	FONA	BBI	BPI	BCa
20	2.762, 4.656	2.651, 4.451	2.933, 4.733	2.933, 4.512	0.2051, 0.2343	0.2082, 0.2371	0.2071, 0.2360	0.2071, 0.2360	0.2246, 0.2507	0.2262, 0.2482	0.2263, 0.2482	0.2263, 0.2482
50	2.988, 4.355	2.890, 4.446	2.909, 4.465	2.939, 4.507	0.2043, 0.3204	0.1820, 0.2961	0.2277, 0.3418	0.2324, 0.3503	0.2514, 0.2649	0.2486, 0.2643	0.2514, 0.2671	0.2525, 0.2686
100	3.182, 3.910	3.201, 3.909	3.113, 3.822	3.124, 3.830	0.2201, 0.2293	0.2196, 0.2292	0.2206, 0.2302	0.2197, 0.2288	0.2412, 0.2603	0.2416, 0.2603	0.2409, 0.2595	0.2400, 0.2589
500	3.536, 3.852	3.529, 3.850	3.543, 3.863	3.543, 3.863	0.2264, 0.2313	0.2263, 0.2312	0.2265, 0.2313	0.2266, 0.2314	0.2504, 0.2575	0.2504, 0.2577	0.2503, 0.2576	0.2498, 0.2572
1000	3.503, 3.741	3.500, 3.749	3.494, 3.743	3.495, 3.743	0.2278, 0.2393	0.2269, 0.2387	0.2284, 0.2402	0.2294, 0.2419	0.2508, 0.2556	0.2508, 0.2555	0.2509, 0.2556	0.2509, 0.2556
5000	3.548, 3.651	3.547, 3.650	3.548, 3.652	3.549, 3.653	0.2259, 0.2277	0.2259, 0.2277	0.2260, 0.2277	0.2261, 0.2279	0.2512, 0.2533	0.2512, 0.2533	0.2512, 0.2533	0.2512, 0.2533
10000	3.633, 3.705	3.633, 3.705	3.633, 3.704	3.633, 3.704	0.2271, 0.2283	0.2271, 0.2282	0.2272, 0.2283	0.2272, 0.2283	0.2519, 0.2534	0.2519, 0.2534	0.2518, 0.2534	0.2518, 0.2534

Where: FONA = First Order Normal Approximation interval; BBI = Basic Bootstrap Interval, BPI = Bootstrap Percentile Interval; BCa = Bias-Corrected Accelerated interval

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From Table 1, the confidence interval for the sample standard error of 10% trimmed mean exhibited a 100% coverage of its estimated sample standard error in all the four approaches and for all the values of B considered in this research work. This same confidence interval for the sample standard error of the 10% trimmed mean was equally getting more precise as the B the number of bootstraps increased. For instance, the estimated sample standard error for the 10% trimmed mean was  $s_{10} = 0.2266338$ ; when B = 20, the bootstrap CI for the standard error returned (0.2051, 0.2343) for the normal approach CI, (0.2082, 0.2371) for the basic, (0.2071, 0.2360) for the percentile, and (0.2071, 0.2360) for the BCa method. As B increased to 5000, it returned (0.2259, 0.2277) for the normal approach, (0.2259, 0.2277) for the basic, (0.2260, 0.2277) for the percentile, and (0.2261, 0.2279) for the BCa method for the same estimated standard error. As B increased further to 10000, the CI for the sample standard error of the 10% trimmed mean returned a more precise interval of (0.2271, 0.2283) for the normal approach, (0.2271, 0.2282) for the basic, (0.2272, 0.2283) for the percentile, and (0.2272, 0.2283) for the BCa showing a shorter interval yet covering the estimate. It returned the best interval for our skewed data set.

The bootstrap CI for  $s_x$  (the sample standard error of the arithmetic mean) covered the estimated value only for B= 20 and 50. As B increased, to 100, the coverage seized for all four approaches and continued like that even up to B = 10000.

At B = 20, the CI for  $s_{20}$  (the standard error of the 20% trimmed mean) did not cover the estimated value. It rendered its best covering as B was increased to 50 and barely covered at B = 100, and 500. Beyond this point, the interval failed to cover the estimate again even for the BCa approach.

No significant difference or improvement in coverage was observed as a result of using a different approach even when B increased to 1000 and even got worse at B=5000 and B = 10000.

It appears that the effect of increasing B should be considered only after the right estimate has been chosen. For instance for skewed data like response data, the interest should be first on choosing the proper estimate like the alpha trimmed mean before one considers the approach to use.

No particular method appears to have specifically managed the effect of the skewness of data better than the other except for the parameter ten percent trimmed mean.

Even though having been adjusted for bias and expected to perform better even skewed data, the result showed that the  $\alpha$ -trimmed mean outperformed it for this set of response data.

We notice that the effect of increasing B was the same irrespective of the approach employed. The difference in its effect was only on the different estimates rather than on the approach employed.

Considering all therefore, we may conclude that for the set of data we considered, the confidence interval for the standard error of the 10% trimmed mean returned the most accurate interval irrespective of the approach employed and the accuracy increased as B increased.

From Table 2, as B the number of bootstrapping increases, we notice that all the four methods of calculating bootstrap CI considered returned shorter confidence intervals. Increasing B however did not improve coverage by any of the methods considered and even for all three statistics of interest.

Table 2: Confidence Interval Length Based on Table 1 above.

B	Standard Error of the Mean ( $s_x$ )				Standard Error of the 10% Trimmed Mean ( $s_{10}$ )				Standard Error of the 20% Trimmed Mean ( $s_{20}$ )			
	FONA	BBI	BPI	BCa	FONA	BBI	BPI	BCa	FONA	BBI	BPI	BCa
20	1.894	1.8	1.8	1.579	0.0292	0.0289	0.0289	0.0289	0.0261	0.022	0.0219	0.0219
50	1.367	1.556	1.556	1.568	0.1161	0.1141	0.1141	0.1179	0.0135	0.0157	0.0157	0.0161
100	0.728	0.708	0.709	0.706	0.0092	0.0096	0.0096	0.0091	0.0191	0.0187	0.0186	0.0189
500	0.316	0.321	0.32	0.32	0.0049	0.0049	0.0048	0.0048	0.0071	0.0073	0.0073	0.0074

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1000	0.238	0.249	0.249	0.248	0.0115	0.0118	0.0118	0.0125	0.0048	0.0047	0.0047	0.0047
5000	0.103	0.103	0.104	0.104	0.0018	0.0018	0.0017	0.0018	0.0021	0.0021	0.0021	0.0021
10000	0.072	0.072	0.071	0.071	0.0012	0.0011	0.0011	0.0011	0.0015	0.0015	0.0016	0.0016

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Note: The meaning of the abbreviations remains the same as in Table 1.

We also observed from Table 2 that for CI of the standard error of the arithmetic mean, the BCa returned the shortest interval among all four approaches when B was as low as 20 returning a confidence interval length of 1.57. However, the confidence interval for the standard error of the arithmetic mean was longer than that for the  $\alpha$  - trimmed mean for the same B=20. This could be because of the fact that the data is skewed. The BCa approach therefore, recommends itself for use when data is skewed and the arithmetic mean is the measure of central tendency. For this skewed data, the confidence interval for the sample  $\alpha$  - trimmed mean performed better even when the B is as low as 20.

Increasing the number of bootstrapping B shortens the interval length but may not guarantee accuracy as it does not guarantee the inclusion of a parameter within the interval except if the right parameter is employed for the right set of data.

#### 4. CONCLUSION: AND RECOMMENDATIONS

##### 4.1 Conclusion

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The 95% confidence interval for the standard error of the 10% trimmed mean returned accurate coverage of the true parameter by all four approaches applied.

In the light of the empirical study, we conclude that increasing B the number of bootstrap samples increases the chances of more precise and shorter confidence intervals rather than increasing the chances for coverage and accuracy of the interval.

##### 5.2 RECOMMENDATIONS

In general, increasing B produces shorter interval rather than increasing coverage/accuracy however, a shorter interval is not a guarantee for better confidence interval unless the probability of coverage is first guaranteed. In the light of the study therefore, we make the following recommendations:

1. In estimating the bootstrap standard error, attention should be more on employing the appropriate measure of central tendency and its sample standard error rather than on arbitrarily increasing B the number of bootstrap samples.
2. When dealing with response data as in the study here, one should seriously consider its skewed nature and as such, the standard error of the alpha trimmed mean should be considered irrespective of the bootstrap method of estimating the CI.
3. Once the right tool is employed, increasing B will enhance the accuracy of the bootstrap CI.
4. With probability of coverage guaranteed, our findings tend to support the opinion of Efron and Tibshirani (1986)[7] and later confirmed by that in estimating the bootstrap confidence interval,

B = 1000 is enough for a well appreciable precise and short interval irrespective of the bootstrap approach employed. B = 500 can render a good short interval really.

5. If the standard error of the arithmetic mean is to be used, the BCa approach to estimating the confidence interval should be considered first.

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